

Martingale approach: Denote by $\mathbb{G} := (\mathcal{G}_n)_{n \in \mathbb{N}_0}$ the filtration generated by $(S_n)_{n \in \mathbb{N}_0}$, i.e. $\mathcal{G}_n = \sigma(S_1, \dots, S_n) = \sigma(Y_1, \dots, Y_n)$, $n \in \mathbb{N}_0$. Note that Y_{n+1} and \mathcal{G}_n are independent for all $n \in \mathbb{N}_0$. Moreover, \tilde{J}_{-1} is a \mathbb{G} -stopping time. Define the process $M = (M_n)_{n \geq 0}$ by

$$M_n := \left(\frac{q}{p}\right)^{S_n}. \quad (7)$$

Then M is a \mathbb{G} -martingale under \mathbb{P}_0 since for $n \in \mathbb{N}_0$ by independence of Y_{n+1} and \mathcal{G}_n we have

$$\begin{aligned} \mathbb{E}_0[M_{n+1} | \mathcal{G}_n] &= [\mathbb{E}_0[M_n \times (q/p)^{Y_{n+1}} | \mathcal{G}_n] = M_n \mathbb{E}_0[(q/p)^{Y_{n+1}}] \\ &= M_n \times \left(p \times \frac{q}{p} + q \times \frac{p}{q}\right) = M_n \quad \mathbb{P}_0\text{-a.s.} \end{aligned} \quad (8)$$

By the *optional stopping theorem*, we have

$$\mathbb{E}_0[M_{\tilde{J}_{-1} \wedge \ell}] = \mathbb{E}_0[M_0] = 1 \quad \text{for all } \ell \in \mathbb{N}. \quad (9)$$

Since $q/p < 1$ for $p > 1/2$, we have $M_{\tilde{J}_{-1} \wedge \ell} \leq (q/p)^{-1} = p/q$ for all $\ell \in \mathbb{N}$. Hence by dominated convergence we have

$$\mathbb{E}_0[M_{\tilde{J}_{-1}} \mathbf{1}_{\{\tilde{J}_{-1} < \infty\}}] = \lim_{\ell \rightarrow \infty} \mathbb{E}_0[M_{\tilde{J}_{-1} \wedge \ell}] = 1. \quad (10)$$

On the other hand we have

$$\mathbb{E}_0[M_{\tilde{J}_{-1}} \mathbf{1}_{\{\tilde{J}_{-1} < \infty\}}] = \mathbb{E}_0[p/q \mathbf{1}_{\{\tilde{J}_{-1} < \infty\}}] = p/q \mathbb{P}_0[\tilde{J}_{-1} < \infty]. \quad (11)$$

This implies that $\mathbb{P}_0[\tilde{J}_{-1} < \infty] = q/p < 1$.

- b) Assume that $(X_n)_{n \in \mathbb{N}}$ is the canonical Markov chain with state space E and transition matrix R . First note that $\alpha = \rho_{1,0} = \rho_{2,1}$ by the special form of R . Using the normal and the strong Markov property we get

$$\begin{aligned} \alpha &= \rho_{1,0} = \mathbb{P}_1[\tilde{H}_0 < \infty] = \mathbb{P}_1[\tilde{H}_0 < \infty, X_1 = 0] + \mathbb{P}_1[\tilde{H}_0 < \infty, X_1 = 2] \\ &= \mathbb{P}_1[X_1 = 0] + \mathbb{P}_1[\tilde{H}_0 \circ \theta_1 < \infty, X_1 = 2] \\ &= q + \mathbb{E}_1[\mathbb{E}_1[\mathbf{1}_{\{\tilde{H}_0 \circ \theta_1 < \infty\}} \mathbf{1}_{\{X_1=2\}} | \mathcal{F}_1]] \\ &= q + \mathbb{E}_1[\mathbf{1}_{\{X_1=2\}} \mathbb{E}_{X_1}[\mathbf{1}_{\{\tilde{H}_0 < \infty\}}]] \\ &= q + p \mathbb{E}_2[\mathbf{1}_{\{\tilde{H}_0 < \infty\}}] = q + p \mathbb{E}_2[\mathbf{1}_{\{\tilde{H}_0 < \infty\}} \mathbf{1}_{\{\tilde{H}_1 < \infty\}}] \\ &= q + p \mathbb{E}_2[\mathbf{1}_{\{\tilde{H}_0 \circ \theta_{\tilde{H}_1} < \infty\}} \mathbf{1}_{\{\tilde{H}_1 < \infty\}}] \\ &= q + p \mathbb{E}_2[\mathbb{E}_2[\mathbf{1}_{\{\tilde{H}_0 \circ \theta_{\tilde{H}_1} < \infty\}} \mathbf{1}_{\{\tilde{H}_1 < \infty\}} | \mathcal{F}_{\tilde{H}_1}]] \\ &= q + p \mathbb{E}_2[\mathbf{1}_{\{\tilde{H}_1 < \infty\}} \mathbb{E}_1[\mathbf{1}_{\{\tilde{H}_0 < \infty\}}]] \\ &= q + p \mathbb{E}_2[\mathbf{1}_{\{\tilde{H}_1 < \infty\}}] \mathbb{E}_1[\mathbf{1}_{\{\tilde{H}_0 < \infty\}}] \\ &= q + p \alpha^2. \end{aligned} \quad (12)$$

The above equation has the solutions $\alpha = 1$ and $\alpha = q/p$. Since, $\alpha \leq 1$ and $q/p \geq 1$ for $p \leq 1/2$, it follows immediately that $\alpha = 1$ for $p \leq 1/2$. Finally, since $\rho_{0,1} = 1$, we have for $q \leq 1/2$

$$\rho_{0,0} \geq \rho_{0,1} \rho_{1,0} = 1, \quad (13)$$

i.e. 0 is a recurrent state.

- (c) Again assume that $(X_n)_{n \in \mathbb{N}}$ is the canonical Markov chain with state space E and transition matrix R . First note that $\beta = \mathbb{E}_1[\tilde{H}_0] = \mathbb{E}_2[\tilde{H}_1]$ by the special form of R . Assuming that $(X_n)_{n \geq 0}$ is positive recurrent and using the normal and the strong Markov property we get

$$\begin{aligned}
\beta &= \mathbb{E}_1[\tilde{H}_0] = \mathbb{E}_1[\tilde{H}_0 \mathbf{1}_{\{X_1=0\}}] + \mathbb{E}_1[\tilde{H}_0 \mathbf{1}_{\{X_1=1\}}] \\
&= \mathbb{E}_1[\mathbf{1}_{\{X_1=0\}}] + \mathbb{E}_1[\tilde{H}_0 \mathbf{1}_{\{X_1=2\}}] \\
&= q + \mathbb{E}_1[(\tilde{H}_0 \circ \theta_1 + 1) \mathbf{1}_{\{X_1=2\}}] \\
&= q + \mathbb{E}_1[\mathbb{E}_1[(\tilde{H}_0 \circ \theta_1 + 1) \mathbf{1}_{\{X_1=2\}} \mid \mathcal{F}_1]] \\
&= q + \mathbb{E}_1[\mathbf{1}_{\{X_1=2\}} \mathbb{E}_{X_1}[\tilde{H}_0 + 1]] \\
&= q + p \mathbb{E}_2[\tilde{H}_0 + 1] = 1 + p \mathbb{E}_2[\tilde{H}_0] \\
&= 1 + p \mathbb{E}_2[\tilde{H}_0 \circ \theta_{\tilde{H}_1} + \tilde{H}_1] \\
&= 1 + p \mathbb{E}_2[\tilde{H}_1] + p \mathbb{E}_2[\mathbb{E}_2[\tilde{H}_0 \circ \theta_{\tilde{H}_1} \mid \mathcal{F}_{\tilde{H}_1}]] \\
&= 1 + p\beta + p \mathbb{E}_2[\mathbb{E}_1[\tilde{H}_0]] = 1 + p\beta + p \mathbb{E}_1[\tilde{H}_0] \\
&= 1 + 2p\beta.
\end{aligned} \tag{14}$$

The above equation has a nonnegative real solution $\beta = 1/(1-2p)$ if and only if $0 < p < 1/2$. Putting this together with the result from part b) this implies that $(X_n)_{n \in \mathbb{N}_0}$ is null recurrent if $p = 1/2$. Finally, we know from an example from the lecture that there exists a stationary distribution for $(X_n)_{n \in \mathbb{N}_0}$ if $0 < p < 1/2$. Thus we may deduce from that $(X_n)_{n \in \mathbb{N}_0}$ is positive recurrent if $0 < p < 1/2$.

- (d) For $p < 1/2$, there exists a reversible distribution and therefore stationary for the reflected random walk. By a theorem of the lecture this stationary distribution π is unique and given by

$$\begin{aligned}
\pi(0) &= \frac{q-p}{2q}, \\
\pi(n) &= \left(\frac{p}{q}\right)^{n-1} \frac{q-p}{2q^2}, \text{ for } n \in \mathbb{N} \setminus \{0\}.
\end{aligned}$$

Solution 10.2

- a) We define for $x \neq y$

$$r'_{x,y} = r_{x,y} \left(\frac{\pi(y)r_{y,x}}{\pi(x)r_{x,y}} \wedge 1 \right)$$

and

$$r'_{x,x} = r_{x,x} + \sum_{y \in E \setminus \{x\}} r_{x,y} \left(1 - \left(\frac{\pi(y)r_{y,x}}{\pi(x)r_{x,y}} \wedge 1 \right) \right).$$

Clearly $r'_{x,y} \geq 0$ for all $x, y \in E$. Furthermore

$$\begin{aligned}
\sum_{y \in E} r'_{x,y} &= \sum_{y \in E \setminus \{x\}} r_{x,y} \left(\frac{\pi(y)r_{y,x}}{\pi(x)r_{x,y}} \wedge 1 \right) + r_{x,x} + \sum_{y \in E \setminus \{x\}} r_{x,y} \left(1 - \left(\frac{\pi(y)r_{y,x}}{\pi(x)r_{x,y}} \wedge 1 \right) \right) \\
&= \sum_{y \in E} r_{x,y} = 1,
\end{aligned}$$

hence $(r'_{x,y})_{x,y \in E}$ is a transition probability.

b) We have to show that for all $x, y \in E$ we have

$$\pi(x)r'_{x,y} = \pi(y)r'_{y,x}.$$

Let $x \neq y \in E$. If $\frac{\pi(y)r_{y,x}}{\pi(x)r_{x,y}} < 1$, then $\frac{\pi(x)r_{x,y}}{\pi(y)r_{y,x}} \geq 1$, hence

$$\pi(x)r'_{x,y} = \pi(x)r_{x,y} \frac{\pi(y)r_{y,x}}{\pi(x)r_{x,y}} = \pi(y)r_{y,x} \underbrace{\left(\frac{\pi(x)r_{x,y}}{\pi(y)r_{y,x}} \wedge 1 \right)}_{=1} = \pi(y)r'_{y,x}.$$

If $\frac{\pi(y)r_{y,x}}{\pi(x)r_{x,y}} \geq 1$, then $\frac{\pi(x)r_{x,y}}{\pi(y)r_{y,x}} < 1$, hence

$$\begin{aligned} \pi(x)r'_{x,y} &= \pi(x)r_{x,y} = \frac{\pi(x)r_{x,y}}{\pi(y)r_{y,x}} \pi(y)r_{y,x} = \pi(y)r'_{y,x}. \\ &= \underbrace{\left(\frac{\pi(x)r_{x,y}}{\pi(y)r_{y,x}} \wedge 1 \right)}_{=1} \pi(y)r_{y,x} \end{aligned}$$

Remark. The Metropolis-Hastings algorithm is useful for generating samples from a distribution π in case only the relative likelihoods $\pi(y)/\pi(x)$ are known (but the normalizing constant is difficult to compute). After a burn-in period, the simulated Markov chain is assumed to be close to the stationary (reversible) distribution π , and samples can be drawn. Care has to be taken due to the serial correlation in this Markov chain, e.g. by only using every n -th sample, where the jump size n is large.

Solution 10.3

(a) Note that we have

$$\nu_i r_{i,j} = \nu_j r_{j,i}.$$

Summing up over i gives

$$\sum_{i \in \mathbb{N}_0} \nu_i r_{i,j} = \sum_{i \in \mathbb{N}_0} \nu_j r_{j,i} = \nu_j \sum_{i \in \mathbb{N}_0} r_{j,i} = \nu_j.$$

(b) Note that the birth and death process is irreducible as all states are connected. Due to the statement proved in Exercise 9.2 (b) it follows from $\nu_i r_{i,j} = \nu_j r_{j,i}$ that the chain is reversible.