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# A Tauberian theorem for the statistical generalized Nörlund-Euler summability method

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Abstract. Let  $(p_n)$  and  $(q_n)$  be any two non-negative real sequences with

$$R_{n} := \sum_{k=0}^{n} p_{k}q_{n-k} \neq 0 \ (n \in \mathbb{N}).$$

With  $\mathsf{E}_n^1-$  we will denote the Euler summability method. Let  $(x_n)$  be a sequence of real or complex numbers and set

$$N_{p,q}^{n} E_{n}^{1} := \frac{1}{R_{n}} \sum_{k=0}^{n} p_{k} q_{n-k} \frac{1}{2^{k}} \sum_{\nu=0}^{k} \binom{k}{\nu} x_{\nu}$$

for  $n \in \mathbb{N}$ . In this paper, we present necessary and sufficient conditions under which the existence of the st- limit of  $(x_n)$  follows from that of  $st - N_{p,q}^n E_n^1$ - limit of  $(x_n)$ . These conditions are one-sided or two-sided if  $(x_n)$  is a sequence of real or complex numbers, respectively.

## 1 Introduction

In what follows we give the concept of the summability method known as the generalized Nörlund summability method (N, p, q) (see [1]). Given two

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non-negative sequences  $(p_n)$  and  $(q_n)$ , the convolution  $(p \star q)$  is defined by

$$R_n := (p \star q)_n = \sum_{k=0}^n p_k q_{n-k} = \sum_{k=0}^n p_{n-k} q_k.$$

In this paper we suppose  $R_n \to \infty$  as  $n \to \infty$ . With  $E_n^1$  – we will denote the Euler summability method. Let  $(x_n)$  be a sequence. When  $(p \star q)_n \neq 0$  for all  $n \in \mathbb{N}$ , the generalized Nörlund-Euler transform of the sequence  $(x_n)$  is the sequence  $N_{p,q}^n E_n^1$  obtained by putting

$$N_{p,q}^{n} E_{n}^{1} = \frac{1}{(p \star q)_{n}} \sum_{k=0}^{n} p_{k} q_{n-k} \frac{1}{2^{k}} \sum_{\nu=0}^{k} \binom{k}{\nu} x_{\nu}.$$
 (1)

We say that the sequence  $(x_n)$  is generalized Nörlund-Euler summable to L determined by the sequences  $(p_n)$  and  $(q_n)$  or briefly summable  $N_{p,q}^n E_n^1$  to L if

$$\lim_{n \to \infty} N_{p,q}^n E_n^1 = L.$$
 (2)

Suppose throughout the paper we assume that the sequence  $q = (q_n)$  satisfies the following conditions:

$$q_{t_n-k} \le 2q_{n-k}, k = 0, 1, 2, 3, \cdots, n; t > 1,$$
(3)

$$q_{n-k} \le 2q_{t_n-k}, k = 0, 1, 2, 3, \cdots, t_n; 0 < t < 1,$$
(4)

where  $t_n = [t \cdot n]$ . If

$$\lim_{n \to \infty} x_n = L \tag{5}$$

implies (2), then the summation method generated by  $N_{p,q}^{n} E_{n}^{1}$  is regular, and it is satisfied under certain conditions. However, the converse is not always true. We can show by the following example

**Example 1** Let us consider that  $x = (x_k) = (-1)^k$ , then we have

$$\left|\frac{1}{R_n}\sum_{k=0}^n p_k q_{n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} (-1)^{\nu}\right| \leq \frac{1}{R_n} \sum_{k=0}^n p_k q_{n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} 1 \to 1 \text{ as } n \to \infty.$$

And as we know  $\mathbf{x} = (\mathbf{x}_k)$ , is not convergent.

Notice that (2) may imply (5) under a certain condition, which is called a Tauberian condition. Any theorem which states that convergence of a sequence follows from its  $N_{p,q}^{n} E_{n}^{1}$  summability and some Tauberian condition is said to be a Tauberian theorem for the  $N_{p,q}^{n} E_{n}^{1}$  summability method. The inclusion and Tauberian type theorems are proved in the papers [6, 7, 2, 3, 4, 5, 8], and some theorems of inclusion, Tauberian and convexity type for certain families of generalized Nörlund methods are obtained in [9].

### 2 Results

In this paper, we present necessary and sufficient conditions under which the existence of the limit  $st - \lim_{n\to\infty} x_n = L$  follows from that of  $st - \lim_{n\to\infty} N_{p,q}^n E_n^1 = L$ . These conditions are one-sided or two-sided if  $(x_n)$  is a sequence of real or complex numbers, respectively.

**Definition 1** A sequence  $(x_n)$  is statistically convergent to L, if for every  $\epsilon > 0$ , we have

$$\lim_{n\to\infty}\frac{|\{k\leq n: |x_k-L|\geq \varepsilon\}|}{n}=0,$$

where |A|, stands for cardinality of the set.

The theory of Tauberian is extensively studied by many authors ([1], [2], [3], [4], [7], [9]). In this section our aim is to find conditions (so-called Tauberian) under which the converse implication holds, for defined convergence. Exactly, we will prove under which conditions convergence of sequences  $(x_n)$ , follows from  $N_{p,q}^n E_n^1$  – convergence.

**Definition 2** A sequence  $(x_n)$  is weighted  $N_{p,q}^n E_n^1$ -statistically convergent to L if for every  $\epsilon > 0$ ,

$$\lim_{n\to\infty}\frac{1}{R_n}\left|\left\{k\leq R_n: \left|\frac{1}{R_n}\sum_{k=0}^n p_k q_{n-k}\frac{1}{2^k}\sum_{\nu=0}^k \binom{k}{\nu}x_\nu - L\right|\geq \varepsilon\right\}\right| = 0.$$

And we say that the sequence  $(x_n)$  is statistically summable to L by the weighted summability method  $N_{p,q}^n E_n^1$ , if  $st - \lim_n N_{p,q}^n E_n^1 = L$ . We denote by  $N_{p,q}^n E_n^1(st)$  the set of all sequences which are statistically summable.

**Theorem 1** If sequence  $x = (x_n)$  is  $N_{p,q}^n E_n^1$  summable to L, then sequence  $x = (x_n)$  is  $N_{p,q}^n E_n^1$  - statistically convergent to L. But not conversely.

**Proof.** The first part of the proof is obvious. To prove the second part we will show this example:

**Example 2** Let us consider that  $(p_k) = (1)$ ,  $(q_{n-k}) = (2^k)$ , and define

$$\mathbf{x}_{\mathbf{k}} = \left\{ egin{array}{ccc} \sqrt{2^{\mathbf{k}}} &, & for \quad \mathbf{k} = 2^{\mathbf{n}} \\ \mathbf{0} &, & otherwise \end{array} 
ight.$$

Under this conditions we get:

$$\frac{1}{2^n-1}\left|\left\{k\leq 2^n-1: \left|\frac{1}{2^n-1}\sum_{k=0}^n 1\cdot \sum_{\nu=0}^k \binom{k}{\nu}x_\nu - 0\right| \geq \varepsilon\right\}\right| \leq \frac{\sqrt{2^n}}{2^n-1} \to 0.$$

Hence, it is  $N_{p,q}^{n}E_{n}^{1}$  statistically summable to 0. On the other hand, if we take in consideration that  $k = 2^{n}$ , we get

$$\frac{1}{2^n-1}\sum_{k=0}^n 1\cdot \sum_{\nu=0}^k \binom{k}{\nu} x_\nu \to \infty, \quad as \quad n \to \infty.$$

From last relation follows that  $\mathbf{x} = (\mathbf{x}_n)$  is not  $N_{p,q}^n E_n^1$  summable to 0.

**Theorem 2** Let us suppose that sequence  $(\boldsymbol{x}_n)\text{-statistically convergent to }\boldsymbol{L},$  and

$$\sup_{\substack{0 \le \nu \le k \\ 0 \le k \le n \\ n \in \mathbb{N}}} (\nu + k + n) |x_{\nu} - L| < \infty.$$

Then it converges  $N_{p,q}^n E_n^1\mbox{-statistically to }L.$  Converse is not true.

**Proof.** From fact that  $(x_n)$  converges statistically to L, we get

$$\lim_{n\to\infty}\frac{|\{k\leq n: |x_k-L|\geq \varepsilon\}|}{n}=0.$$

Let us denote by  $B_{\varepsilon} = \{k \leq n : |x_k - L| \geq \varepsilon\}$  and  $\overline{B_{\varepsilon}} = \{k \leq n : |x_k - L| < \varepsilon\}$ . Then

$$\left|\frac{1}{R_n}\sum_{k=0}^n p_k q_{n-k}\frac{1}{2^k}\sum_{\nu=0}^k \binom{k}{\nu} x_{\nu} - L\right| = \left|\frac{1}{R_n}\sum_{k=0}^n p_k q_{n-k}\frac{1}{2^k}\sum_{\nu=0}^k \binom{k}{\nu} (x_{\nu} - L)\right| \leq \frac{1}{R_n}\sum_{k=0}^n p_k q_{n-k}\frac{1}{2^k}\sum_{\nu=0}^k \binom{k}{\nu} (x_{\nu} - L)\right| \leq \frac{1}{R_n}\sum_{k=0}^n p_k q_{n-k}\frac{1}{2^k}\sum_{\nu=0}^k \binom{k}{\nu} (x_{\nu} - L)$$

$$\frac{1}{R_n}\sum_{\substack{k=0\\k\in B_\varepsilon}}^n p_k q_{n-k}\frac{1}{2^k}\sum_{\nu=0}^k \binom{k}{\nu}|x_\nu-L| + \frac{1}{R_n}\sum_{\substack{k=0\\k\in \overline{B}_\varepsilon}}^n p_k q_{n-k}\frac{1}{2^k}\sum_{\nu=0}^k \binom{k}{\nu}|x_\nu-L| \leq \frac{1}{R_n}\sum_{\substack{k=0\\k\in \overline{B}_\varepsilon}}^n p_k q_{n-k}\frac{1}{2^k}\sum_{\nu=0}^k \binom{k}{\nu} p_k$$

(from given conditions for sequence  $(x_\nu),$  there exists a constant C such that  $|x_\nu-L|\leq \frac{C}{\nu+k+n})$ 

$$\begin{split} \frac{C}{R_n} \sum_{\substack{k=0\\k\in B_\varepsilon}}^n p_k q_{n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} \frac{1}{\nu+k+n} + \frac{\varepsilon}{R_n} \sum_{\substack{k=0\\k\in \overline{B_\varepsilon}}}^n p_k q_{n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} \le \\ \frac{C}{R_n} \sum_{\substack{k=0\\k\in B_\varepsilon}}^n p_k q_{n-k} \frac{1}{2^k} \frac{1}{k+n} \sum_{\nu=0}^k \binom{k}{\nu} + \frac{\varepsilon}{R_n} \sum_{\substack{k=0\\k\in \overline{B_\varepsilon}}}^n p_k q_{n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} \le \\ \le \frac{C|B_\varepsilon|}{n} \cdot \frac{\max_{0\le k\le n} \{p_k q_{n-k}\}}{R_n} + \varepsilon \to 0 + \varepsilon, \quad \text{as} \quad n \to \infty. \end{split}$$

To show that converse is not true we will use into consideration this

**Example 3** Let us consider that  $(p_n) = 1$ ,  $(q_{n-k}) = (2^k)$  for  $n \in \mathbb{N}$  and we define the sequence  $x = (x_n)$ , as follows:

$$x_k = \left\{ \begin{array}{ll} 1 & , \quad \mathrm{for} \quad k=m^2-m, \cdots, m^2-1 \\ -m & , \quad \mathrm{for} \quad k=m^2, m=2, \cdots \\ 0 & , \quad \mathrm{otherwise} \end{array} \right.$$

Under this conditions, after some calculations we get:

$$\left|\frac{1}{2^{n}-1}\sum_{k=0}^{n}1\cdot\sum_{\nu=0}^{k}\binom{k}{\nu}x_{\nu}-1\right| \leq \left|\frac{1}{2^{n}-1}\sum_{k=0}^{n}1\cdot\sum_{\nu=0}^{k}\binom{k}{\nu}-1\right| = 0$$

From last relation follows that  $\mathbf{x} = (\mathbf{x}_n)$  is  $N_{p,q}^n E_n^1$  – summable to 1. Hence from Theorem 1,  $(\mathbf{x}_n)$  is  $N_{p,q}^n E_n^1$  – statistically convergent. On the other hand, the sequence  $(\mathbf{m}^2; \mathbf{m} = 2, 3 \cdots,)$  has natural density zero and it is clear that  $\mathbf{st} - \liminf_n \mathbf{x}_n = 0$  and  $\mathbf{st} - \limsup_n \mathbf{x}_n = 1$ . Thus,  $(\mathbf{x}_k)$  is not statistically convergent.

Theorem 3 If

$$st - \liminf_{n} \frac{R_{t_n}}{R_n} > 1, t > 1$$
(6)

where  $t_n$ , denotes the integral parts of the [tn] for every  $n \in \mathbb{N}$ , and let  $(x_k)$  be a sequence of real numbers which converges to L,  $N_{p,q}^n E_n^1$  – statistically. Then  $(x_k)$  is st – convergent to the same number L if and only if the following two conditions hold:

$$\inf_{t>1} \limsup_{n} \frac{1}{R_{n}} \left| \left\{ k \le R_{n} : \frac{1}{R_{t_{k}} - R_{k}} \sum_{j=k+1}^{t_{k}} p_{j} q_{t_{k}-j} \frac{1}{2^{j}} \sum_{\nu=0}^{j} {j \choose \nu} (x_{\nu} - x_{k}) \le -\epsilon \right\} \right| = 0$$
(7)

and

$$\inf_{0 < t < 1} \limsup_{n} \sup \frac{1}{R_{n}} \left| \left\{ k \le R_{n} : \frac{1}{R_{k} - R_{t_{k}}} \sum_{j=t_{k}+1}^{k} p_{j} q_{k-j} \frac{1}{2^{j}} \sum_{\nu=0}^{j} {j \choose \nu} (x_{k} - x_{\nu}) \le -\epsilon \right\} \right| = 0.$$
(8)

**Remark 1** Let us suppose that  $st - \lim_k x_k = L$ ;  $(x_n)$  is  $N_{p,q}^n E_n^1$  – statistically convergent and relation (6) satisfies, then for every t > 1, is valid the following relation:

$$st - \lim_{k} \frac{1}{R_{t_k} - R_k} \sum_{j=k+1}^{t_k} p_j q_{t_k - j} \frac{1}{2^j} \sum_{\nu=0}^{j} {j \choose \nu} (x_{\nu} - x_k) = 0$$
(9)

and in case where  $0 < t < 1,\,$ 

$$st - \lim_{k} \frac{1}{R_k - R_{t_k}} \sum_{j=t_k+1}^{k} p_j q_{k-j} \frac{1}{2^j} \sum_{\nu=0}^{j} {j \choose \nu} (x_k - x_{\nu}) = 0.$$
(10)

In the next result, we will consider the case where  $x = (x_n)$  is a sequence of complex numbers.

**Theorem 4** Let us suppose that relation (6) is satisfied. And  $(x_n)$  be a sequence of complex numbers, which is  $N_{p,q}^n E_n^1$  – statistically convergent to L. Then  $(x_n)$  is st – convergent to the same number L if and only if the following two conditions hold:

$$\inf_{t>1} \limsup_{n} \sup \frac{1}{R_n} \left| \left\{ k \le R_n : \left| \frac{1}{R_{t_k} - R_k} \sum_{j=k+1}^{t_k} p_j q_{t_k-j} \frac{1}{2^j} \sum_{\nu=0}^j \binom{j}{\nu} (x_\nu - x_k) \right| \ge \varepsilon \right\} \right| = 0$$

$$(11)$$

and

$$\inf_{0 < t < 1} \limsup_{n} \frac{1}{R_n} \left| \left\{ k \le R_n : \left| \frac{1}{R_k - R_{t_k}} \sum_{j=t_k+1}^k p_j q_{k-j} \frac{1}{2^j} \sum_{\nu=0}^j \binom{j}{\nu} (x_k - x_\nu) \right| \ge \varepsilon \right\} \right| = 0.$$

$$(12)$$

In what follows we will show some auxiliary lemmas which are needful in the sequel.

**Lemma 1** Condition given by relation (6) is equivalent to this one:

$$st - \lim_{n} \inf \frac{R_n}{R_{t_n}} > 1, \quad 0 < t < 1.$$
(13)

**Proof.** Let us suppose that relation (6) is valid, 0 < t < 1 and  $m = t_n = [t \cdot n]$ ,  $n \in \mathbb{N}$ . Then it follows that

$$\frac{1}{t}>1 \Rightarrow \frac{m}{t}=\frac{[t\cdot n]}{t}\leq n,$$

from above relation we obtain:

$$\frac{R_n}{R_{t_n}} \geq \frac{R_{[\frac{m}{t}]}}{R_{t_n}} \Rightarrow st - \liminf_n \inf \frac{R_n}{R_{t_n}} \geq st - \liminf_n \frac{R_{[\frac{m}{t}]}}{R_{t_n}} > 1.$$

Conversely, let us suppose that relation (13) is valid. Let t > 1 be given number and let  $t_1$  be chosen such that  $1 < t_1 < t$ . Set  $m = t_n = [t \cdot n]$ . From  $0 < \frac{1}{t} < \frac{1}{t_1} < 1$ , it follows that:

$$\mathfrak{n} \leq \frac{\mathfrak{t}\mathfrak{n}-\mathfrak{l}}{\mathfrak{t}_1} < \frac{[\mathfrak{t}\mathfrak{n}]}{\mathfrak{t}_1} = \frac{\mathfrak{m}}{\mathfrak{t}_1},$$

provided  $t_1 \leq t - \frac{1}{n}$ , which is a case where if n is large enough. Under this conditions we have:

$$\frac{R_{t_n}}{R_n} \geq \frac{R_{t_n}}{R_{\left\lfloor\frac{m}{t_1}\right\rfloor}} \Rightarrow st_{\lambda} - \liminf_{n} \frac{R_{t_n}}{R_n} \geq st_{\lambda} - \liminf_{n} \frac{R_{t_n}}{R_{\left\lfloor\frac{m}{t_1}\right\rfloor}} > 1.$$

**Lemma 2** Let us suppose that relation (6) is satisfied and let  $x = (x_k)$  be a sequence of complex numbers which is  $N_{p,q}^n E_n^1$ -statistically convergent to L. Then for every t > 0,

$$st - \lim_n N_{p,q}^{t_n} E_n^1 = L.$$

**Proof.** (I) Let us consider that t > 1. Then

$$\lim_{n \to \infty} \frac{1}{R_n} \sum_{k=0}^n p_k q_{n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} x_{\nu} = \lim_{n \to \infty} \frac{1}{R_{t_n}} \sum_{k=0}^{t_n} p_k q_{t_n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} x_{\nu}, \quad (14)$$

and for every  $\epsilon > 0$  we have:

$$\{k \le R_{t_n} : |N_{p,q}^{t_n} E_n^1 - L| \ge \varepsilon\} \subset \{k \le R_n : |N_{p,q}^n E_n^1 - L| \ge \varepsilon\} \cup$$

$$\left\{k \le R_n : \frac{1}{R_n} \sum_{k=0}^n p_k q_{n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} x_{\nu} \neq \frac{1}{R_{t_n}} \sum_{k=0}^{t_n} p_k q_{t_n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} x_{\nu} \right\}.$$

Now proof of the lemma in this case follows from relation (14) and  $st - \lim_{n \to \infty} N_{p,q}^n E_n^1 = L$ .

(II) In this case we have that 0 < t < 1. For  $t_n = [t \cdot n]$ , for any natural number n, we can conclude that  $N_{p,q}^{t_n} E_n^1$  does not appears more than  $[1 + t^{-1}]$  times in the sequence  $N_{p,q}^n E_n^1$ . In fact if there exist integers k, l such that

$$n \leq t \cdot k < t(k+1) < \cdots < t(k+l-1) < n+1 \leq t(k+l),$$

then

$$n + t(l-1) \le t(k+l-1) < n+1 \Rightarrow l < 1 + \frac{l}{t}.$$

And we have this estimation

$$\begin{split} &\frac{1}{R_n} \left| \left\{ k \leq R_n : |N_{p,q}^{t_n} E_n^1 - L| \geq \varepsilon \right\} \right| \\ &\leq \left( 1 + \frac{1}{t} \right) \frac{1}{R_n} \left| \left\{ k \leq R_{t_n} : |N_{p,q}^n E_n^1 - L| \geq \varepsilon \right\} \right| \\ &\leq 2(1+t) \frac{1}{R_{t_n}} \left| \left\{ k \leq R_{t_n} : |N_{p,q}^n E_n^1 - L| \geq \varepsilon \right\} \right|, \end{split}$$

provided  $\frac{1}{R_n}(\frac{t+1}{t}) \leq 2(t+1)\frac{1}{R_{t_n}}$ , which is the case where n is large enough. From last relation it follows:  $st - \lim_n N_{p,q}^{t_n} E_n^1 = L$ .

**Proposition 1** Let us suppose that relation (6) is satisfied and let  $x = (x_k)$  be a sequence of complex numbers which is  $N_{p,q}^{t_n} E_n^1$ -statistically convergent to L. Then for every t > 1,

$$st - \lim_{k} \frac{1}{R_{t_{k}} - R_{k}} \sum_{j=k+1}^{t_{k}} p_{j} q_{t_{k}-j} \frac{1}{2^{j}} \sum_{\nu=0}^{j} {j \choose \nu} x_{\nu} = L;$$
(15)

and for every 0 < t < 1,

$$st - \lim_{k} \frac{1}{R_{k} - R_{t_{k}}} \sum_{j=t_{k}+1}^{k} p_{j} q_{k-j} \frac{1}{2^{j}} \sum_{\nu=0}^{j} {j \choose \nu} x_{\nu} = L.$$
(16)

**Proof.** (I) Let us consider the case where t > 1. Then we obtain

$$\begin{split} &\frac{1}{R_{t_n} - R_n} \sum_{k=n+1}^{t_n} p_k q_{t_n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} (x_\nu - L) \\ &= \frac{R_{t_n}}{R_{t_n} - R_n} \frac{1}{R_{t_n}} \sum_{k=0}^{t_n} p_k q_{t_n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} (x_\nu - L) \\ &- \frac{R_n}{R_{t_n} - R_n} \frac{1}{R_n} \sum_{k=0}^n p_k q_{n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} (x_\nu - L) \\ &= \frac{R_{t_n}}{R_{t_n} - R_n} \frac{1}{R_{t_n}} \sum_{k=0}^{t_n} p_k q_{t_n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} (x_\nu - L) \\ &- \frac{R_{t_n}}{R_{t_n} - R_n} \frac{1}{R_{t_n}} \sum_{k=0}^n p_k q_{t_n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} (x_\nu - L) \\ &= \frac{R_{t_n}}{R_{t_n} - R_n} \frac{1}{R_{t_n}} \sum_{k=0}^n p_k (q_{t_n-k} + q_{n-k} - q_{n-k}) \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} (x_\nu - L) \\ &= \frac{R_n}{R_{t_n} - R_n} \frac{1}{R_{t_n}} \sum_{k=0}^n p_k q_{n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} (x_\nu - L) \\ &- \frac{R_n}{R_{t_n} - R_n} \frac{1}{R_n} \sum_{k=0}^n p_k q_{n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} (x_\nu - L) \\ &- \frac{R_n}{R_{t_n} - R_n} \frac{1}{R_n} \sum_{k=0}^n p_k (q_{t_n-k} - q_{n-k}) \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} (x_\nu - L). \end{split}$$

From

$$\limsup_{n} \sup \frac{\mathsf{R}_{\mathsf{t}_{n}}}{\mathsf{R}_{\mathsf{t}_{n}} - \mathsf{R}_{n}} < \infty, \tag{18}$$

definition of the sequence  $(q_n)$ , Lemma 2 and relation (17), we get relation (15).

(II) In this case we have that 0 < t < 1. Then

$$\frac{1}{R_n - R_{t_n}} \sum_{k=t_n+1}^n p_k q_{n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} x_{\nu} = \frac{R_n}{R_n - R_{t_n}} \frac{1}{R_n} \sum_{k=0}^n p_k q_{n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} x_{\nu} - \frac{1}{2^k} \sum_{\nu=0}^k \sum_{\nu=0}^k \frac{1}{2^k} \sum_{\nu=0}^k \frac{1}{2^k} \sum_{\nu=0}^k \sum_{\nu=0}^k \sum_{\nu=0}^k \sum_{\nu=0}^k \sum_{\nu=0}^k \sum_$$

$$\begin{aligned} \frac{R_{t_n}}{R_n - R_{t_n}} \frac{1}{R_{t_n}} \sum_{k=0}^{t_n} p_k q_{n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} x_\nu &= \frac{R_{t_n}}{R_n - R_{t_n}} \frac{1}{R_{t_n}} \sum_{k=0}^n p_k q_{n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} x_\nu - \frac{1}{R_n - R_{t_n}} \sum_{k=0}^n p_k (q_{n-k} - q_{t_n-k}) \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} x_\nu. \end{aligned}$$
Now proof of the proposition is similar to the first part.

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#### 3 Proofs of the theorems

**Proof of Theorem 3.** Necessity. Suppose that  $\lim_{n\to\infty} x_n = L$ , and (6) holds. Following Proposition 1, we have

$$\begin{split} st &- \lim_{n \to \infty} \frac{1}{R_{\lambda_n} - R_n} \sum_{k=n+1}^{\lambda_n} p_k q_{\lambda_n - k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} (x_\nu - x_n) = \\ st &- \lim_{n \to \infty} \frac{1}{R_{\lambda_n} - R_n} \sum_{k=n+1}^{\lambda_n} p_k q_{\lambda_n - k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} x_\nu - x_n = 0, \end{split}$$

for every  $\lambda > 1$ . In case where  $0 < \lambda < 1$ , we find that

$$\begin{split} st &- \lim_{n \to \infty} \frac{1}{R_n - R_{\lambda_n}} \sum_{k=\lambda_n+1}^n p_k q_{n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} (x_n - x_\nu) = \\ x_n - st &- \lim_{n \to \infty} \frac{1}{R_n - R_{\lambda_n}} \sum_{k=\lambda_n+1}^n p_k q_{n-k} \frac{1}{2^k} \sum_{\nu=0}^k \binom{k}{\nu} x_\nu = 0. \end{split}$$

Sufficiency. Assume that conditions (7) and (8) are satisfied. In what follows we will prove that  $st-\lim_{n\to\infty} x_n = L$ . Or equivalently,  $st-\lim (N_{p,q}^n E_n^1 - x_n) =$ 0. First we will consider the case where t > 1. We will start from this estimation:

$$\begin{split} x_{n} &- N_{p,q}^{n} E_{n}^{1} \\ &= \frac{R_{t_{n}}}{R_{t_{n}} - R_{n}} \left[ \frac{1}{R_{t_{n}}} \sum_{j=0}^{t_{n}} p_{j} q_{t_{n}-j} \frac{1}{2^{j}} \sum_{\nu=0}^{j} {j \choose \nu} x_{\nu} - \frac{1}{R_{n}} \sum_{j=0}^{n} p_{j} q_{n-j} \frac{1}{2^{j}} \sum_{\nu=0}^{j} {j \choose \nu} x_{\nu} \right] \\ &- \frac{1}{R_{t_{n}} - R_{n}} \sum_{j=n+1}^{t_{n}} p_{j} q_{t_{n}-j} \frac{1}{2^{j}} \sum_{\nu=0}^{j} {j \choose \nu} (x_{\nu} - x_{n}). \end{split}$$

For any  $\epsilon > 0$ , we obtain:

$$\begin{split} \{k \leq R_n : x_k - N_{p,q}^n E_n^1 \geq \varepsilon\} \subset \left\{k \leq R_n : \frac{R_{t_k}}{R_{t_k} - R_k} (N_{p,q}^{t_k} E_k^1 - N_{p,q}^k E_k^1) \geq \frac{\varepsilon}{2}\right\} \cup \\ \left\{k \leq R_n : \frac{1}{R_{t_k} - R_k} \sum_{j=k+1}^{t_k} p_j q_{t_k-j} \frac{1}{2^j} \sum_{\nu=0}^j \binom{j}{\nu} (x_\nu - x_k) \leq -\frac{\varepsilon}{2}\right\}. \end{split}$$

From relation (7), it follows that for every  $\gamma > 0$ , exists a t > 1 such that

$$\lim_{n} \sup \frac{1}{R_{n}} \left| \left\{ k \leq R_{n} : \frac{1}{R_{t_{k}} - R_{k}} \sum_{j=k+1}^{t_{k}} p_{j} q_{t_{k}-j} \frac{1}{2^{j}} \sum_{\nu=0}^{j} {j \choose \nu} (x_{\nu} - x_{k}) \leq -\varepsilon \right\} \right| \leq \gamma.$$

By Lemma 2 and relation (18) we get

$$\lim_{n} \sup \frac{1}{R_{n}} \left| \left\{ k \leq R_{n} : |R_{t_{k}}(R_{t_{k}} - R_{k})^{-1}(N_{p,q}^{t_{k}}E_{k}^{1} - N_{p,q}^{k}E_{k}^{1})| \geq \frac{\varepsilon}{2} \right\} \right| = 0.$$

Combining last three relations we have:

$$\lim_{n} \sup \frac{1}{R_{n}} \left| \left\{ k \leq R_{n} : x_{k} - N_{p,q}^{k} E_{k}^{1} \geq \epsilon \right\} \right| \leq \gamma,$$

and  $\gamma$  is arbitrary, we conclude that for every  $\varepsilon > 0$ ,

$$\lim_{n} \sup \frac{1}{R_{n}} \left| \left\{ k \le R_{n} : x_{k} - N_{p,q}^{k} E_{k}^{1} \ge \epsilon \right\} \right| = 0.$$
<sup>(19)</sup>

Now we consider case where 0 < t < 1. From above we get that:

$$\begin{split} x_{n} &- N_{p,q}^{n} E_{n}^{1} \\ &= \frac{R_{t_{n}}}{R_{n} - R_{t_{n}}} \left[ \frac{1}{R_{n}} \sum_{j=0}^{n} p_{j} q_{n-j} \frac{1}{2^{j}} \sum_{\nu=0}^{j} {j \choose \nu} x_{\nu} - \frac{1}{R_{t_{n}}} \sum_{j=0}^{t_{n}} p_{j} q_{t_{n}-j} \frac{1}{2^{j}} \sum_{\nu=0}^{j} {j \choose \nu} x_{\nu} \right] \\ &+ \frac{1}{R_{n} - R_{t_{n}}} \sum_{j=t_{n}+1}^{n} p_{j} q_{n-j} \frac{1}{2^{j}} \sum_{\nu=0}^{j} {j \choose \nu} (x_{n} - x_{\nu}). \end{split}$$

For any  $\varepsilon > 0$ ,

$$\{k \leq R_n : x_k - N_{p,q}^k E_k^1 \geq \varepsilon\} \subset \left\{k \leq R_n : \frac{R_{t_k}}{R_k - R_{t_k}} (N_{p,q}^k E_k^1 - N_{p,q}^{t_k} E_k^1) \geq \frac{\varepsilon}{2}\right\} \cup$$

$$\left\{k \leq R_n: \frac{1}{R_k - R_{t_k}} \sum_{j=t_k+1}^k p_j q_{k-j} \frac{1}{2^j} \sum_{\nu=0}^j \binom{j}{\nu} (x_k - x_\nu) \leq -\frac{\varepsilon}{2}\right\}.$$

For same reasons as in the case where t > 1, by Lemma 2, we have that for every  $\epsilon > 0$ ,

$$\limsup_{n} \frac{1}{R_n} \left| \left\{ k \le R_n : x_k - N_{p,q}^k E_k^1 \le -\epsilon \right\} \right| = 0.$$
 (20)

Finally from relations (19) and (20) we get:

$$\limsup_{n} \frac{1}{R_{n}} \left| \left\{ k \leq R_{n} : |x_{k} - N_{p,q}^{k} E_{k}^{1}| \geq \varepsilon \right\} \right| = 0. \quad \Box$$

**Remark 2** Let us suppose that  $st - \lim_k x_k = L$ ,  $st - \lim_k N_{p,q}^k E_k^1 = L$  and relation (6) satisfies. Then for every t > 1, relation (9) holds, and in case where 0 < t < 1, relation (10) is valid.

**Proof of Theorem 4.** We omit it because it is similar to the Theorem 3.  $\Box$ 

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