

## Research Article

# On the Strong Laws for Weighted Sums of $\rho^*$ -Mixing Random Variables

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Complete convergence is studied for linear statistics that are weighted sums of identically distributed  $\rho^*$ -mixing random variables under a suitable moment condition. The results obtained generalize and complement some earlier results. A Marcinkiewicz-Zygmund-type strong law is also obtained.

## 1. Introduction

Suppose that  $\{X_n; n \geq 1\}$  is a sequence of random variables and  $S$  is a subset of the natural number set  $N$ . Let  $F_S = \sigma(X_i; i \in S)$ ,

$$\rho_n^* = \sup \left\{ \text{corr}(f, g) : \forall S \times T \subset N \times N, \text{dist}(S, T) \geq n, \forall f \in L^2(F_S), g \in L^2(F_T) \right\}, \quad (1.1)$$

where

$$\text{corr}(f, g) = \frac{\text{Cov}\{f(X_i; i \in S), g(X_j; j \in T)\}}{[\text{Var}\{f(X_i; i \in S)\} \text{Var}\{g(X_j; j \in T)\}]^{1/2}}. \quad (1.2)$$

*Definition 1.1.* A random variable sequence  $\{X_n; n \geq 1\}$  is said to be a  $\rho^*$ -mixing random variable sequence if there exists  $k \in N$  such that  $\rho_k^* < 1$ .

The notion of  $\rho^*$ -mixing seems to be similar to the notion of  $\rho$ -mixing, but they are quite different from each other. Many useful results have been obtained for  $\rho^*$ -mixing random variables. For example, Bradley [1] has established the central limit theorem, Byrc and Smoleński [2] and Yang [3] have obtained moment inequalities and the strong law of large numbers, Wu [4, 5], Peligrad and Gut [6], and Gan [7] have studied almost sure convergence, Utev and Peligrad [8] have established maximal inequalities and the invariance principle, An and Yuan [9] have considered the complete convergence and Marcinkiewicz-Zygmund-type strong law of large numbers, and Budsaba et al. [10] have proved the rate of convergence and strong law of large numbers for partial sums of moving average processes based on  $\rho^-$ -mixing random variables under some moment conditions.

For a sequence  $\{X_n; n \geq 1\}$  of i.i.d. random variables, Baum and Katz [11] proved the following well-known complete convergence theorem: suppose that  $\{X_n; n \geq 1\}$  is a sequence of i.i.d. random variables. Then  $EX_1 = 0$  and  $E|X_1|^{rp} < \infty$  ( $1 \leq p < 2, r \geq 1$ ) if and only if  $\sum_{n=1}^{\infty} n^{r-2} P(|\sum_{i=1}^n X_i| > n^{1/p} \varepsilon) < \infty$  for all  $\varepsilon > 0$ .

Hsu and Robbins [12] and Erdős [13] proved the case  $r = 2$  and  $p = 1$  of the above theorem. The case  $r = 1$  and  $p = 1$  of the above theorem was proved by Spitzer [14]. An and Yuan [9] studied the weighted sums of identically distributed  $\rho^*$ -mixing sequence and have the following results.

**Theorem B.** *Let  $\{X_n; n \geq 1\}$  be a  $\rho^*$ -mixing sequence of identically distributed random variables,  $\alpha p > 1$ ,  $\alpha > 1/2$ , and suppose that  $EX_1 = 0$  for  $\alpha \leq 1$ . Assume that  $\{a_{ni}; 1 \leq i \leq n\}$  is an array of real numbers satisfying*

$$\sum_{i=1}^n |a_{ni}|^p = O(\delta), \quad 0 < \delta < 1, \quad (1.3)$$

$$\#A_{nk} = \#\{1 \leq i \leq n : |a_{ni}|^p > (k+1)^{-1}\} \geq ne^{-1/k}. \quad (1.4)$$

If  $E|X_1|^p < \infty$ , then

$$\sum_{n=1}^{\infty} n^{\alpha p - 2} P\left(\max_{1 \leq j \leq n} \left| \sum_{i=1}^j a_{ni} X_i \right| > \varepsilon n^{\alpha}\right) < \infty. \quad (1.5)$$

**Theorem C.** *Let  $\{X_n; n \geq 1\}$  be a  $\rho^*$ -mixing sequence of identically distributed random variables,  $\alpha p > 1$ ,  $\alpha > 1/2$ , and  $EX_1 = 0$  for  $\alpha \leq 1$ . Assume that  $\{a_{ni}; 1 \leq i \leq n\}$  is array of real numbers satisfying (1.3). Then*

$$n^{-1/p} \sum_{i=1}^n a_{ni} X_i \longrightarrow 0 \text{ a.s. } (n \longrightarrow \infty). \quad (1.6)$$

Recently, Sung [15] obtained the following complete convergence results for weighted sums of identically distributed NA random variables.

**Theorem D.** Let  $\{X, X_n; n \geq 1\}$  be a sequence of identically distributed NA random variables, and let  $\{a_{ni}; 1 \leq i \leq n, n \geq 1\}$  be an array of constants satisfying

$$A_\alpha = \limsup_{n \rightarrow \infty} A_{\alpha,n} < \infty, \quad A_{\alpha,n} = \sum_{i=1}^n \frac{|a_{ni}|^\alpha}{n} \quad (1.7)$$

for some  $0 < \alpha \leq 2$ . Let  $b_n = n^{1/\alpha}(\log n)^{1/\gamma}$  for some  $\gamma > 0$ . Furthermore, suppose that  $EX = 0$  where  $1 < \alpha \leq 2$ . If

$$\begin{aligned} E|X|^\alpha &< \infty, \quad \text{for } \alpha > \gamma, \\ E|X|^\alpha \log|X| &< \infty, \quad \text{for } \alpha = \gamma, \\ E|X|^\gamma &< \infty, \quad \text{for } \alpha < \gamma, \end{aligned} \quad (1.8)$$

then

$$\sum_{n=1}^{\infty} \frac{1}{n} P\left(\max_{1 \leq j \leq n} \left| \sum_{i=1}^j a_{ni} X_i \right| > b_n \varepsilon\right) < \infty \quad \forall \varepsilon > 0. \quad (1.9)$$

We find that the proof of Theorem C is mistakenly based on the fact that (1.5) holds for  $\alpha p = 1$ . Hence, the Marcinkiewicz-Zygmund-type strong laws for  $\rho^*$ -mixing sequence have not been established.

In this paper, we shall not only partially generalize Theorem D to  $\rho^*$ -mixing case, but also extend Theorem B to the case  $\alpha p = 1$ . The main purpose is to establish the Marcinkiewicz-Zygmund strong laws for linear statistics of  $\rho^*$ -mixing random variables under some suitable conditions.

We have the following results.

**Theorem 1.2.** Let  $\{X, X_n; n \geq 1\}$  be a sequence of identically distributed  $\rho^*$ -mixing random variables, and let  $\{a_{ni}; 1 \leq i \leq n, n \geq 1\}$  be an array of constants satisfying

$$A_\beta = \limsup_{n \rightarrow \infty} A_{\beta,n} < \infty, \quad A_{\beta,n} = \sum_{i=1}^n \frac{|a_{ni}|^\beta}{n}, \quad (1.10)$$

where  $\beta = \max(\alpha, \gamma)$  for some  $0 < \alpha \leq 2$  and  $\gamma > 0$ . Let  $b_n = n^{1/\alpha}(\log n)^{1/\gamma}$ . If  $EX = 0$  for  $1 < \alpha \leq 2$  and (1.8) for  $\alpha \neq \gamma$ , then (1.9) holds.

*Remark 1.3.* The proof of Theorem D was based on Theorem 1 of Chen et al. [16], which gave sufficient conditions about complete convergence for NA random variables. So far, it is not known whether the result of Chen et al. [16] holds for  $\rho^*$ -mixing sequence. Hence, we use different methods from those of Sung [15]. We only extend the case  $\alpha \neq \gamma$  of Theorem D to  $\rho^*$ -mixing random variables. It is still open question whether the result of Theorem D about the case  $\alpha = \gamma$  holds for  $\rho^*$ -mixing sequence.

**Theorem 1.4.** *Under the conditions of Theorem 1.2, the assumptions  $EX = 0$  for  $1 < \alpha \leq 2$  and (1.8) for  $\alpha \neq \gamma$  imply the following Marcinkiewicz-Zygmund strong law:*

$$b_n^{-1} \sum_{i=1}^n a_{ni} X_i \longrightarrow 0 \text{ a.s. } (n \longrightarrow \infty). \quad (1.11)$$

## 2. Proof of the Main Result

Throughout this paper, the symbol  $C$  represents a positive constant though its value may change from one appearance to next. It proves convenient to define  $\log x = \max(1, \ln x)$ , where  $\ln x$  denotes the natural logarithm.

To obtain our results, the following lemmas are needed.

**Lemma 2.1** (Utev and Peligrad [8]). *Suppose  $N$  is a positive integer,  $0 \leq r < 1$ , and  $q \geq 2$ . Then there exists a positive constant  $D = D(N, r, q)$  such that the following statement holds.*

*If  $\{X_i; i \geq 1\}$  is a sequence of random variables such that  $\rho_N^* \leq r$  with  $EX_i = 0$  and  $E|X_i|^q < \infty$  for every  $i \geq 1$ , then for all  $n \geq 1$ ,*

$$E \left( \max_{1 \leq i \leq n} |S_i|^q \right) \leq D \left( \sum_{i=1}^n E|X_i|^q + \left( \sum_{i=1}^n EX_i^2 \right)^{q/2} \right), \quad (2.1)$$

where  $S_i = \sum_{j=1}^i X_j$ .

**Lemma 2.2.** *Let  $X$  be a random variable and  $\{a_{ni}; 1 \leq i \leq n, n \geq 1\}$  be an array of constants satisfying (1.10),  $b_n = n^{1/\alpha} (\log n)^{1/\gamma}$ . Then*

$$\sum_{n=1}^{\infty} n^{-1} \sum_{i=1}^n P(|a_{ni} X| > b_n) \leq \begin{cases} CE|X|^\alpha & \text{for } \alpha > \gamma, \\ CE|X|^\gamma & \text{for } \alpha < \gamma. \end{cases} \quad (2.2)$$

*Proof.* If  $\gamma > \alpha$ , by  $\sum_{i=1}^n |a_{ni}|^\gamma = O(n)$  and Lyapounov's inequality, then

$$\frac{1}{n} \sum_{i=1}^n |a_{ni}|^\alpha \leq \left( \frac{1}{n} \sum_{i=1}^n |a_{ni}|^\gamma \right)^{\alpha/\gamma} = O(1). \quad (2.3)$$

Hence, (1.7) is satisfied. From the proof of (2.1) of Sung [15], we obtain easily that the result holds.  $\square$

*Proof of Theorem 1.2.* Let  $X_{ni} = a_{ni}X_i I(|a_{ni}X_i| \leq b_n)$ . For all  $\varepsilon > 0$ , we have

$$\begin{aligned} \sum_{n=1}^{\infty} \frac{1}{n} P\left(\max_{1 \leq j \leq n} \left| \sum_{i=1}^j a_{ni} X_i \right| > \varepsilon b_n\right) &\leq \sum_{n=1}^{\infty} \frac{1}{n} P\left(\max_{1 \leq j \leq n} |a_{nj} X_j| > b_n\right) + \sum_{n=1}^{\infty} \frac{1}{n} P\left(\max_{1 \leq j \leq n} \left| \sum_{i=1}^j X_{ni} \right| > \varepsilon b_n\right) \\ &:= I_1 + I_2. \end{aligned} \quad (2.4)$$

To obtain (1.9), we need only to prove that  $I_1 < \infty$  and  $I_2 < \infty$ .

By Lemma 2.2, one gets

$$I_1 \leq \sum_{n=1}^{\infty} \frac{1}{n} \sum_{j=1}^n P(|a_{nj} X_j| > b_n) = \sum_{n=1}^{\infty} \frac{1}{n} \sum_{j=1}^n P(|a_{nj} X| > b_n) < \infty. \quad (2.5)$$

Before the proof of  $I_2 < \infty$ , we prove firstly

$$b_n^{-1} \max_{1 \leq j \leq n} \left| \sum_{i=1}^j E a_{ni} X_i I(|a_{ni} X_i| \leq b_n) \right| \rightarrow 0, \quad \text{as } n \rightarrow \infty. \quad (2.6)$$

For  $0 < \alpha \leq 1$ ,

$$\begin{aligned} b_n^{-1} \max_{1 \leq j \leq n} \left| \sum_{i=1}^j E a_{ni} X_i I(|a_{ni} X_i| \leq b_n) \right| &\leq b_n^{-1} \sum_{i=1}^n E |a_{ni} X_i| I(|a_{ni} X_i| \leq b_n) \leq b_n^{-\alpha} \sum_{i=1}^n |a_{ni}|^{\alpha} E |X|^{\alpha} \\ &\leq C(\log n)^{-\alpha/\gamma} E |X|^{\alpha} \rightarrow 0, \quad \text{as } n \rightarrow \infty. \end{aligned} \quad (2.7)$$

For  $1 < \alpha \leq 2$ ,

$$\begin{aligned} b_n^{-1} \max_{1 \leq j \leq n} \left| \sum_{i=1}^j E a_{ni} X_i I(|a_{ni} X_i| \leq b_n) \right| &= b_n^{-1} \max_{1 \leq j \leq n} \left| \sum_{i=1}^j E a_{ni} X_i I(|a_{ni} X_i| > b_n) \right| (EX_i = 0) \\ &\leq b_n^{-1} \sum_{i=1}^n E |a_{ni} X_i| I(|a_{ni} X_i| > b_n) \leq b_n^{-\alpha} \sum_{i=1}^n |a_{ni}|^{\alpha} E |X|^{\alpha} \\ &\leq C(\log n)^{-\alpha/\gamma} E |X|^{\alpha} \rightarrow 0, \quad \text{as } n \rightarrow \infty. \end{aligned} \quad (2.8)$$

Thus (2.6) holds. So, to prove  $I_2 < \infty$ , it is enough to show that

$$I_3 = \sum_{n=1}^{\infty} \frac{1}{n} P\left(\max_{1 \leq j \leq n} \left| \sum_{i=1}^j X_{ni} - EX_{ni} \right| > \varepsilon b_n\right) < \infty, \quad \forall \varepsilon > 0. \quad (2.9)$$

By the Chebyshev inequality and Lemma 2.1, for  $q \geq \max\{2, \gamma\}$ , we have

$$\begin{aligned}
 I_3 &\leq C \sum_{n=1}^{\infty} n^{-1} b_n^{-q} E \left( \max_{1 \leq j \leq n} \left| \sum_{i=1}^j X_{ni} - EX_{ni} \right|^q \right) \\
 &\leq C \sum_{n=1}^{\infty} n^{-1} b_n^{-q} \sum_{i=1}^n E |a_{ni} X_i|^q I(|a_{ni} X_i| \leq b_n) \\
 &\quad + C \sum_{n=1}^{\infty} n^{-1} b_n^{-q} \left[ \sum_{i=1}^n E (a_{ni} X_i)^2 I(|a_{ni} X_i| \leq b_n) \right]^{q/2} \\
 &:= I_{31} + I_{32}.
 \end{aligned} \tag{2.10}$$

For  $I_{31}$ , we consider the following two cases.

If  $\alpha < \gamma$ , note that  $E|X|^\gamma < \infty$ . We have

$$I_{31} \leq C \sum_{n=1}^{\infty} n^{-1} b_n^{-\gamma} \sum_{i=1}^n |a_{ni}|^\gamma E|X|^\gamma \leq C \sum_{n=1}^{\infty} n^{-\frac{\gamma}{\alpha}} (\log n)^{-1} < \infty. \tag{2.11}$$

If  $\alpha > \gamma$ , note that  $E|X|^\alpha < \infty$ . we have

$$I_{31} \leq C \sum_{n=1}^{\infty} n^{-1} b_n^{-\alpha} \sum_{i=1}^n |a_{ni}|^\alpha E|X|^\alpha \leq C \sum_{n=1}^{\infty} n^{-1} (\log n)^{-\alpha/\gamma} < \infty. \tag{2.12}$$

Next, we prove  $I_{32} < \infty$  in the following two cases.

If  $\alpha < \gamma \leq 2$  or  $\gamma < \alpha \leq 2$ , take  $q > \max(2, 2\gamma/\alpha)$ . Noting that  $E|X|^\alpha < \infty$ , we have

$$\begin{aligned}
 I_{32} &\leq C \sum_{n=1}^{\infty} n^{-1} b_n^{-\alpha q/2} \left[ \sum_{i=1}^n |a_{ni}|^\alpha E|X|^\alpha \right]^{q/2} \\
 &\leq C \sum_{n=1}^{\infty} n^{-1} (\log n)^{-\alpha q/(2\gamma)} < \infty.
 \end{aligned} \tag{2.13}$$

If  $\gamma > 2 \geq \alpha$  or  $\gamma \geq 2 > \alpha$ , one gets  $E|X|^2 < \infty$ . Since  $\sum_{i=1}^n |a_{ni}|^\alpha = O(n)$ , it implies  $\max_{1 \leq i \leq n} |a_{ni}|^\alpha \leq Cn$ . Therefore, we have

$$\sum_{i=1}^n |a_{ni}|^k = \sum_{i=1}^n |a_{ni}|^\alpha |a_{ni}|^{k-\alpha} \leq C n n^{(k-\alpha)/\alpha} = C n^{k/\alpha} \tag{2.14}$$

for all  $k \geq \alpha$ . Hence,  $\sum_{i=1}^n |a_{ni}|^2 = O(n^{2/\alpha})$ . Taking  $q > \gamma$ , we have

$$\begin{aligned} I_{32} &\leq C \sum_{n=1}^{\infty} n^{-1} b_n^{-q} \left[ \sum_{i=1}^n |a_{ni}|^2 \right]^{q/2} \\ &\leq C \sum_{n=1}^{\infty} n^{-1} b_n^{-q} n^{q/\alpha} = C \sum_{n=1}^{\infty} n^{-1} (\log n)^{-q/\gamma} < \infty. \end{aligned} \quad (2.15)$$

□

*Proof of Theorem 1.4.* By (1.9), a standard computation (see page 120 of Baum and Katz [11] or page 1472 of An and Yuan [9]), and the Borel-Cantelli Lemma, we have

$$\frac{\max_{1 \leq j \leq 2^i} \left| \sum_{i=1}^j a_{ni} X_i \right|}{2^{(i+1)/\alpha} (\log 2^{i+1})^{1/\gamma}} \rightarrow 0 \text{ a.s. } (i \rightarrow \infty). \quad (2.16)$$

For any  $n \geq 1$ , there exists an integer  $i$  such that  $2^{i-1} \leq n < 2^i$ . So

$$\max_{2^{i-1} \leq n < 2^i} \frac{\left| \sum_{j=1}^n a_{nj} X_j \right|}{b_n} \leq \frac{\max_{1 \leq j \leq 2^i} \left| \sum_{i=1}^j a_{nj} X_j \right|}{2^{(i-1)/\alpha} (\log 2^{i-1})^{1/\gamma}} = 2^{2/\alpha} \frac{\max_{1 \leq j \leq 2^i} \left| \sum_{j=1}^n a_{nj} X_j \right|}{2^{(i+1)/\alpha} (\log 2^{i+1})^{1/\gamma}} \left( \frac{i+1}{i-1} \right)^{1/\gamma}. \quad (2.17)$$

From (2.16) and (2.17), we have

$$\lim_{n \rightarrow \infty} b_n^{-1} \sum_{i=1}^n a_{ni} X_i = 0 \text{ a.s.} \quad (2.18)$$

□

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