

Applied Stochastic Processes

Solution Sheet 2

Solution 2.1

- (a) Let A be a bounded Borel set. Notice that for some $n_0 \in \mathbb{N}$,

$$A \subset [0, a_n], \quad \forall n \geq n_0.$$

Let us compute the characteristic function of $N_n(A)$ for $n \geq n_0$ at $t \in \mathbb{R}$:

$$\begin{aligned} \phi_{N_n(A)}(t) &= \mathbb{E} \left[e^{itN_n(A)} \right] \\ &= \prod_{i=1}^n \mathbb{E} \left[e^{it1(S_i \in A)} \right] \\ &= \mathbb{E} \left[e^{it1(S_1 \in A)} \right]^n \\ &= \left(\left(1 - \frac{|A|}{a_n} \right) + \frac{|A|}{a_n} e^{it} \right)^n \\ &= \left(1 + \frac{\lambda|A|}{n} (e^{it} - 1) \right)^n \\ &\xrightarrow{n \rightarrow \infty} \exp(\lambda|A| (e^{it} - 1)) \end{aligned}$$

The function $t \mapsto \exp(\lambda|A| (e^{it} - 1))$ is continuous at 0, so the sequence $(N_n(A))_{n \in \mathbb{N}}$ converges in distribution towards a Poisson distributed random variable with parameter $\lambda|A|$.

- (b) Let $(t_1, \dots, t_k) \in \mathbb{R}^k$, and A_1, A_2, \dots, A_k be Borel sets such that for $n \geq n_0$, we have $A_1, A_2, \dots, A_k \subset [0, a_n]$. We compute the characteristic function of the random vector $(N^n(A_1), N^n(A_2), \dots, N^n(A_k))$ at point $(t_1, \dots, t_k) \in \mathbb{R}^k$:

$$\begin{aligned}
\phi_{(N^n(A_1), N^n(A_2), \dots, N^n(A_k))}(t_1, \dots, t_k) &= \mathbb{E} \left[e^{i \sum_{j=1}^k t_j N^n(A_j)} \right] \\
&= \mathbb{E} \left[\exp \left(i \sum_{j=1}^k t_j \sum_{l=1}^n 1(S_l \in A_j) \right) \right] \\
&= \mathbb{E} \left[\exp \left(i \sum_{l=1}^n \sum_{j=1}^k t_j 1(S_l \in A_j) \right) \right] \\
&= \prod_{l=1}^n \mathbb{E} \left[\exp \left(i \sum_{j=1}^k t_j 1(S_l \in A_j) \right) \right] \\
&= \mathbb{E} \left[\exp \left(i \sum_{j=1}^k t_j 1(X_1 \in A_j) \right) \right]^n \\
&= \left(\left(1 - \frac{\lambda}{n} \sum_{j=1}^k |A_j| \right) + \sum_{j=1}^k e^{it_j} \frac{\lambda |A_j|}{n} \right)^n \\
&\xrightarrow{n \rightarrow \infty} \exp \left(\lambda \left(\sum_{j=1}^k |A_j| (e^{it_j} - 1) \right) \right) \\
&= \prod_{i=1}^k \exp(\lambda (|A_i| (e^{it_i} - 1))),
\end{aligned}$$

where we used for the fourth equality that the X_i 's are independent, for the fifth equality, that the X_i 's are identically distributed, and for the sixth equality, that the X_i 's are uniformly distributed in $[0, \frac{n}{\lambda}]$. The sequence of characteristic functions of the vectors $(N^n(A_1), N^n(A_2), \dots, N^n(A_k))$ converges pointwise towards the product of characteristic functions of Poisson random variables with parameter $\lambda|A_1|, \lambda|A_2|, \dots, \lambda|A_k|$. The limit function is continuous at 0, so the sequence of random vectors converges to a vector of independent Poisson-distributed random variables with parameters $\lambda|A_1|, \lambda|A_2|, \dots, \lambda|A_k|$.

- (c) By the previous question, for $0 \leq t_1 < t_2 < \dots < t_k < \infty$, the sequence of vectors $\left((N_{t_1}^n, N_{t_2}^n - N_{t_1}^n, N_{t_3}^n - N_{t_2}^n, \dots, N_{t_k}^n - N_{t_{k-1}}^n) \right)_{n \in \mathbb{N}}$ converges in distribution towards $((N_{t_1}, N_{t_2} - N_{t_1}, N_{t_3} - N_{t_2}, \dots, N_{t_k} - N_{t_{k-1}}))$ where N is a Poisson process with rate λ . The map that transforms (x_1, x_2, \dots, x_k) into $(x_1, x_2 + x_1, x_3 + x_2 + x_1, \dots, \sum_{j=1}^k x_j)$ is a continuous bijection, therefore the sequence of vectors $((N_{t_1}^n, N_{t_2}^n, N_{t_3}^n, \dots, N_{t_k}^n))_{n \in \mathbb{N}}$ converges in distribution to $(N_{t_1}, N_{t_2}, N_{t_3}, \dots, N_{t_k})$. The processes N^n converge to a Poisson process with rate λ in the sense of finite-dimensional distributions.

Solution 2.2

For $n \in \mathbb{N}$ set $\tilde{T}_n := -\log(U_n)/\lambda$. Clearly, the \tilde{T}_n are i.i.d. as the U_n and we have for $t \geq 0$

$$M_t = \sup \left\{ n \in \mathbb{N}_0 : \sum_{k=1}^n \tilde{T}_k \leq t \right\}. \quad (1)$$

Moreover, the \tilde{T}_n are exponentially distributed. Indeed, let $x \in \mathbb{R}$. Then we have

$$\mathbb{P}[\tilde{T}_1 \leq x] = \mathbb{P}[\log U_1 \geq -\lambda x] = \mathbb{P}[U_1 \geq e^{-\lambda x}] = \begin{cases} 0 & \text{if } x \leq 0, \\ 1 - e^{-\lambda x} & \text{if } x > 0. \end{cases} \quad (2)$$

- a) First, note that for $0 \leq s < t$ we have $\{M_s = +\infty\} \subset \{M_t = +\infty\}$. Hence, we have $\bigcup_{t \geq 0} \{M_t = +\infty\} = \bigcup_{j \in \mathbb{N}} \{M_j = +\infty\}$. To establish the first claim, it therefore suffices to show that for all $j \in \mathbb{N}$ we have $\mathbb{P}[M_j = +\infty] = 0$. Fix $j \in \mathbb{N}$. Using the independence of the \tilde{T}_n we have

$$\begin{aligned} \mathbb{P}[M_j = +\infty] &\leq \mathbb{P}\left[\bigcap_{k \in \mathbb{N}} \{\tilde{T}_k \leq j\}\right] = \prod_{k=1}^{\infty} \mathbb{P}[\tilde{T}_k \leq j] \\ &= \prod_{k=1}^{\infty} (1 - e^{-\lambda j}) = 0. \end{aligned} \quad (3)$$

Next, it follows immediately from the definition that (N_t) starts at 0 and has nondecreasing trajectories with values in \mathbb{N}_0 . It remains to show that the sample paths are right-continuous. Clearly we have to check this property only outside the set $\bigcup_{t \geq 0} \{M_t = +\infty\}$. Fix $t \geq 0$ and let $\omega \in \bigcap_{t \geq 0} \{M_t < \infty\}$ and $n := N_t(\omega) = M_t(\omega) \in \mathbb{N}_0$. Then we have by definition of M_t

$$\sum_{k=1}^n \tilde{T}_k(\omega) \leq t \quad \text{and} \quad \sum_{k=1}^{n+1} \tilde{T}_k(\omega) > t. \quad (4)$$

Hence, for all $\epsilon > 0$ sufficiently small we also have

$$\sum_{k=1}^n \tilde{T}_k(\omega) \leq t + \epsilon \quad \text{and} \quad \sum_{k=1}^{n+1} \tilde{T}_k(\omega) > t + \epsilon. \quad (5)$$

Therefore, for all $\epsilon > 0$ sufficiently small we have $N_{t+\epsilon}(\omega) = n = N_t(\omega)$ implying that the function $s \mapsto N_s(\omega)$ is right-continuous at $s = t$.

- b) Denote by $(S_n)_{n \in \mathbb{N}}$ the sequence of jump times of N . For $\omega \in \bigcap_{t \geq 0} \{M_t < +\infty\}$, it follows immediately from the definition of M and N that $N_t(\omega)$ increase by jumps of size 1 and that $S_n(\omega) = \sum_{k=1}^n \tilde{T}_k(\omega) < \infty$, $n \in \mathbb{N}$. (Note that $\tilde{T}_n(\omega) \in (0, \infty)$ for all $n \in \mathbb{N}$). For $\omega \in \bigcup_{t \geq 0} \{M_t = +\infty\}$, we have $N_t(\omega) \equiv 0$ and $S_n(\omega) = +\infty$ for all $n \in \mathbb{N}$. In conclusion, N increases by jumps of size 1, and we have $S_n < \infty$ \mathbb{P} -a.s. for all $n \in \mathbb{N}$. Denote by $(T_n)_{n \in \mathbb{N}}$ the sequence of interarrival times of N . This is well defined on $\bigcap_{t \geq 0} \{M_t < +\infty\}$, where we have $T_n = \tilde{T}_n$. In particular, the T_n are i.i.d. and distributed as the \tilde{T}_n , i.e. $T_n \sim \text{Exp}(\lambda)$. This establishes the claim as we know from the lecture that a counting process with jumps of size 1 starting at 0 and having i.i.d. interarrival times that are exponentially distributed with parameter $\lambda > 0$, is a Poisson process with rate λ .

Solution 2.3

Denote by $\text{Sym}(n)$ the symmetric group of degree n . Since the X_i have a density, we have $X_{(1)} < X_{(2)} < \dots < X_{(n)}$ \mathbb{P} -a.s. Using that the X_i are i.i.d. and that the order of $\text{Sym}(n)$ is $n!$,

we get for all $B \in \mathcal{B}(\mathbb{R}^n)$

$$\begin{aligned}
\mathbb{P}[(X_{(1)}, \dots, X_{(n)}) \in B] &= \mathbb{P}[(X_{(1)}, \dots, X_{(n)}) \in B, X_{(1)} < X_{(2)} < \dots < X_{(n)}] \\
&= \sum_{\pi \in \text{Sym}(n)} \mathbb{P}[(X_{(1)}, \dots, X_{(n)}) \in B, X_{(1)} < \dots < X_{(n)}, X_{(1)} = X_{\pi(1)}, \dots, X_{(n)} = X_{\pi(n)}] \\
&= \sum_{\pi \in \text{Sym}(n)} \mathbb{P}[(X_{\pi(1)}, \dots, X_{\pi(n)}) \in B, X_{\pi(1)} < X_{\pi(2)} < \dots < X_{\pi(n)}] \\
&= n! \mathbb{P}[(X_1, \dots, X_n) \in B, X_1 < X_2 < \dots < X_n] \\
&= \int_{\mathbb{R}^n} \mathbb{1}_{\{(x_1, \dots, x_n) \in B\}} n! \mathbb{1}_{\{x_1 < x_2 < \dots < x_n\}} \prod_{i=1}^n f(x_i) dx_1 \dots dx_n. \tag{6}
\end{aligned}$$

This establishes the claim.