

$\Rightarrow \left\{ R \in \mathcal{L}_2(Y, X) \Rightarrow \text{Cov}(RZ) \in \mathcal{L}_1(X) \text{ for any } Y\text{-HSP } Z \right\}$
 \uparrow
 constraint for reconstruction

$$T^+ g = \sum_j \sigma_j^{-1} (g_1, g_2)_Y f_j \text{ with sing. system } \{\sigma_j, f_j, g_j\} \text{ for } T$$

General spectral reconstruction

$$Rg = \sum_j \underbrace{\lambda_j}_{\text{weights}} \frac{1}{\sigma_j} (g_1, g_2) f_j \in \mathcal{L}_2, \text{ if } \sum_j \left(\lambda_j / \sigma_j \right)^2 < \infty$$

$\rightarrow \lambda_j$ has to decay much faster than σ_j

e.g. Tikhonov

$$Rg = \sum_j \frac{1}{\alpha + \sigma_j^2} (T^* g, f_j) f_j = \sum_j \frac{\sigma_j}{\alpha + \sigma_j^2} (g, g_j) f_j$$

$$R \in \mathcal{L}_2 \text{ if } \sum_j \left(\frac{\sigma_j}{\alpha + \sigma_j^2} \right)^2 < \infty, \text{ which is satisfied if } T \in \mathcal{L}_2(X, Y)$$

5.1.3 The minimax linear estimator

\rightarrow Setting from above we could also write $\mathcal{L}_2(Y, X)$

Def 5.1.3.B Risk of estimator $R \in \mathcal{L}_2(Y, X)$ under a priori knowledge $f^+ \in K, K \subset X$, and noise model Z

$$\text{risk}_R(Z, K) := \sup_{f \in K} \| R(Tf + Z) - f \|_X^2$$

Implicit assumption: $\text{Cov}(RZ) \in \mathcal{L}_1(X)$

$$\leadsto \| R(Tf + Z) - f \|_X^2 = \mathbb{E}(\| R(Tf + Z) - f \|_X^2)$$

Bias variance decomposition

$$\mathbb{E}(\| R(Tf + Z) - f \|_X^2) = \| (RT - \text{Id})f \|_X^2 + \mathbb{E}(\| RZ \|_X^2) \quad (1.5.2.F)$$

Def 5.1.3.C: minimax linear risk

$$\text{mex}(Z, K) = \inf_{\substack{R \in \mathcal{L}_2(Y, X) \\ \text{Cov}(RZ) \in \mathcal{L}_1}} \text{risk}_R(Z, K)$$

Special: (i) $Z = \varepsilon W, \mathbb{E}(W) = 0, \text{Cov}(W) = \text{Id}$
 \uparrow stochastic noise level

$$(ii) K = K_{\nu, \sigma} := \left\{ f \in X : \sum \sigma_j^{-2\nu} (f, f_j)^2 \leq \nu^2 \right\}$$

$$R_{\{\lambda_j\}} g = \sum_{j=1}^{\infty} \frac{\lambda_j}{\sigma_j} (g - g_j) f_j \in \mathcal{L}_2(Y, X) \quad (5.1.3.D)$$

$$\Rightarrow \text{risk}_{R_{\{\lambda_j\}}}(\varepsilon W, K_{V,S}) = \sup_{f \in K_{V,S}} \| (RT - Id) f \|_X^2 + \varepsilon^2 \| RW \|^2 =$$

$$\| (RT - Id) f \|^2 = \sum_j (\lambda_j - 1)^2 (f, f_j)^2$$

$$= \sup_{f \in K_{V,S}} \sum_j (\lambda_j - 1)^2 (f, f_j)^2 + \varepsilon^2 \text{tr}(RR^*)$$

$$= \sum_j (\lambda_j / \sigma_j)^2$$

$$= \sup_{\substack{x_j^2 \leq \sigma_j^2 \\ \sum_j x_j^2 \leq \rho^2}} \sum_j (1 - \lambda_j)^2 x_j^2 + \varepsilon^2 \sum_j (\lambda_j / \sigma_j)^2$$

$$x_j = \frac{\rho}{\sigma_j} \sigma_j^{-1} x_j$$

$$= \sup_{\sum_j x_j^2 \leq 1} \sum_j \left((1 - \lambda_j)^2 \sigma_j^{2\nu} \rho^2 x_j^2 + \varepsilon^2 (\lambda_j / \sigma_j)^2 \right)$$

$$\text{risk}_{R_{\{\lambda_j\}}}(\varepsilon W, K_{V,S}) = \rho^2 \max_j \left\{ (1 - \lambda_j)^2 \sigma_j^{2\nu} \right\} + \varepsilon^2 \sum_{j=1}^{\infty} (\lambda_j / \sigma_j)^2 \quad (5.1.3.F)$$

Lemma 5.1.3.E If $Z = \varepsilon W$, then every minimax linear estimator (i.e., one that realizes mer) has the form (5.1.3.D)

Proof: $S \in \mathcal{L}_2(Y, X)$, Idea: Build $R_{\{\lambda_j\}}$, which is at least as good as S

Ingenious choice: $\lambda_j := \sigma_j (S g_j, f_j)_X$

Observation: let k be the index that yields the max in (5.1.3.F) for the given sequence $\{\lambda_j\}$

$f^* = \rho \sigma_k^{-\nu} f_k$ is the $f \in K_{V,S}$ that realizes the sup in $\text{risk}_{R_{\{\lambda_j\}}}(\varepsilon W, K_{V,S})$

$$\text{risk}_{R_{\{\lambda_j\}}}(\varepsilon W, K_{V,S}) = \text{risk}_{R_{\{\lambda_j\}}}(\varepsilon W, \{f^*\}) =$$

$$= \underbrace{\| (R_{\{\lambda_j\}} T - Id) f^* \|_X^2}_{= (\lambda_k - 1)^2 \rho^2 \sigma_k^{2\nu}} + \underbrace{\varepsilon^2 (\| R_{\{\lambda_j\}} Z \|^2)}_{= \varepsilon^2 \text{tr}(R_{\{\lambda_j\}}^* R_{\{\lambda_j\}})} \quad (*)$$

$$= (\lambda_k - 1)^2 \rho^2 \sigma_k^{2\nu} + \varepsilon^2 \text{tr}(R_{\{\lambda_j\}}^* R_{\{\lambda_j\}})$$

$$\Gamma(S_T - Id) f^* = \underbrace{\sigma_k^{\nu+1} S S g_k}_{T f_k = \sigma_k g_k} - f^*$$

$$\begin{aligned} \|(S_T - Id) f^*\|_X^2 &= \sum_{j \neq k} \sigma_k^{2\nu+2} S^2 (S g_k, f_j)^2 + \sigma_k^{\nu+1} S (S g_k, f_k - S \sigma_k^\nu)^2 \\ &\geq \underline{S^2 \sigma_k^{2\nu} (\lambda_k - 1)^2} \end{aligned}$$

$$\begin{aligned} (*) \quad \leq \|(S_T - Id) f^*\|_X^2 + \varepsilon^2 \underbrace{\|R_{\lambda_j}\|_{S^2}^2} &= \sum_j \|R_{\lambda_j} g_j\|_X^2 = \\ &= \sum_j (\lambda_j / \sigma_j)^2 = \sum_j (S g_j, f_j)_X^2 = \|S\|_{\mathcal{L}_2}^2 \\ &= \text{tr}(\text{Cov}(S W)) \end{aligned}$$

$$\begin{aligned} \Rightarrow \text{risk}_{R_{\lambda_j}}(\varepsilon W, K_{\nu, S}) &\leq \|(S_T - Id) f^*\|_X^2 + \varepsilon^2 \text{tr}(\text{Cov}(S W)) \\ &= E(\|S(T f^* + z) - f^*\|_X^2) \\ &\leq \text{risk}_S(\varepsilon W, K_{\nu, S}) \end{aligned}$$

□

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$$\text{Recall: } \text{risk}_{R_{\lambda_j}}(\varepsilon W, K_{\nu, S}) = \underbrace{S^2 \sup_j (1 - \lambda_j)^2 \sigma_j^{2\nu}}_{\sup_{f \in K_{\nu, S}} \|(R_{\lambda_j} T - Id) f\|_X^2} + \varepsilon^2 \underbrace{\sum_j (\lambda_j / \sigma_j)^2}_{\text{tr}(\text{Cov}(R_{\lambda_j} z))}$$

Goal: Find λ_j st risk becomes minimal
(For-the-sake-of-simplicity assumption: T injective)

(i) for any minimizing sequence: $0 \leq \lambda_j \leq 1$
 $\Rightarrow \sup_j (1 - \lambda_j)^2 \sigma_j^{2\nu} \leq \sigma_1^{2\nu}$

(ii) $S_k = \left\{ (\lambda_j)_{j \in \mathbb{N}} : 0 \leq \lambda_j \leq 1, \sup_j |1 - \lambda_j| \sigma_j^\nu = k \right\}, 0 \leq k \leq \sigma_1^\nu$

Best sequence in S_k : $\lambda_j^k := \max\{1 - k \sigma_j^{-\nu}, 0\}$

$[(1 - \lambda_j) \sigma_j^\nu = k \rightarrow \lambda_j = 1 - k \sigma_j^{-\nu}$ (must not be < 0 !)]

$\text{risk}_{R_{\lambda_j^k}}(\varepsilon W, K_{\nu, S}) = S^2 k^2 + \varepsilon^2 \sum_{j=1}^{\bar{J}(k)} \left(\frac{\lambda_j^k}{\sigma_j}\right)^2$ (only finitely many λ_j^k are non-zero) (*)