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# Inverse problems

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### The deconvolution problem

Suppose that we have nonnegative observations  $Z_1, \dots, Z_n$  from a distribution with density

$$h_0(z) = \int g(z - x) dF_0(x), \quad z \geq 0,$$

where  $g$  is a known decreasing continuous density on  $[0, \infty)$  and  $F_0$  is the distribution function we want to estimate.

$F_0$  has support, contained in  $[0, \infty)$  (i.e., corresponds to nonnegative random variables).

The maximum likelihood estimator (MLE)  $\hat{F}_n$  of  $F_0$  is obtained by maximizing the **log likelihood**

$$\sum_{i=1}^n \log \int g(Z_i - x) dF(x),$$

over all distribution functions  $F$ .

Conjecture in part 2 of Groeneboom and Wellner (1992): at an interior point  $t$  of the support of  $F_0$ :

$$n^{1/3} \left\{ \hat{F}_n(t) - F_0(t) \right\} \longrightarrow cZ,$$

where  $Z$  is the location of the minimum of 2-sided Brownian motion plus a parabola.

An important step in proving the conjectured behavior in Groeneboom and Wellner (1992) is to write the functional

$$\int_{x \in [0, t)} g(t - x) d\hat{F}_n(x)$$

in the form

$$g(0)\hat{F}_n(t) - \int_{x \in [0, t)} \{g(t - x) - g(0)\} d\hat{F}_n(x), \quad (1)$$

and to show that a centered version of  $\int_{x \in [0, t)} \{g(t - x) - g(0)\} d\hat{F}_n(x)$  is of lower order than  $g(0)\hat{F}_n(t)$ :  $\int_{x \in [0, t)} \{g(t - x) - g(0)\} d\hat{F}_n(x)$  is a so-called **smooth functional**. Note that

$$g(0)\hat{F}_n(t) = \int_{x \in [0, t)} g(0) d\hat{F}_n(x),$$

and that the only crucial difference of the latter integral with the integral in (1) is that the integrand of the integral in (1) is **continuous** at  $x = t$ . We want in fact to prove that

$$\int_{x \in [0, t)} \{g(t - x) - g(0)\} d\hat{F}_n(x) - \int_{x \in [0, t)} \{g(t - x) - g(0)\} dF_0(x) = O_p \left( n^{-1/2} \right),$$

whereas  $\hat{F}_n(t)$  itself will have the so-called “cube root” behavior.

## Integral equations

**Canonical approach:** consider the functional

$$K_t(F) = \int_{x \in [0, t)} \{g(t - x) - g(0)\} dF(x),$$

and let  $a_{t,F}$  solve the (**adjoint**, see below) equation

$$\begin{aligned} [L_F^*(a)](x) &= E \{a_{t,F}(Z) \mid X = x\} \\ &= \int_{z \geq x} a_{t,F}(z) g(z - x) dz = \{g(t - x) - g(0)\} 1_{[0, t)}(x) - K_t(F), \end{aligned} \quad (2)$$

where  $a_{t,F}$  has to be in the range of the **score operator**:

$$a_{t,F}(z) = [L_F(b)](z) = E_F \{b_{t,F}(X) \mid X + Y = z\} = \frac{\int_{[0, z]} b_{t,F}(x) g(z - x) dF(x)}{h_F(z)}. \quad (3)$$

If we could solve these equations for  $\hat{F}_n$ , we would have a representation of the following form:

$$K_t(\hat{F}_n) - K_t(F_0) = \int a_{t, \hat{F}_n}(z) d(H_n - H_0)(z).$$

“Argument”:

$$\begin{aligned} \int a_{t,\hat{F}_n}(z) dH_n(z) &= \int_{x \in [0,\infty)} \frac{\int_{x \in [0,z]} g(z-x) b_{t,\hat{F}_n}(x) d\hat{F}_n(x)}{h_{\hat{F}_n}(z)} dH_n(z) \\ &= \int_{x \in [0,\infty)} \int_{z \geq x} \frac{g(z-x)}{\hat{h}_n(z)} dH_n(z) b_{t,\hat{F}_n}(x) d\hat{F}_n(x) = \int_{x \in [0,\infty)} b_{t,\hat{F}_n}(x) d\hat{F}_n(x) = 0, \end{aligned}$$

and

$$\begin{aligned} \int a_{t,\hat{F}_n}(z) dH_0(z) &= \int a_{t,\hat{F}_n}(z) \int_0^z g(z-x) dF_0(x) dz \\ &= \int_{x \in [0,\infty)} \int_{z \geq x} a_{t,\hat{F}_n}(z) g(z-x) dz dF_0(x) \\ &= \int_{x \in [0,\infty)} \left\{ \{g(z-x) - g(0)\} 1_{[0,t)}(x) - K_t(\hat{F}_n) \right\} dF_0(x) \\ &= \int_{x \in [0,\infty)} \{g(z-x) - g(0)\} 1_{[0,t)}(x) dF_0(x) - K_t(\hat{F}_n) \\ &= K_t(F_0) - K_t(\hat{F}_n). \end{aligned}$$

Unfortunately, there is generally **no**  $b_{t,\hat{F}_n}$  such that

$$a_{t,\hat{F}_n}(z) = E_{\hat{F}_n} \left\{ b_{t,\hat{F}_n}(X) \mid X + Y = z \right\} = \frac{\int_{[0,z]} b_{t,\hat{F}_n}(x) g(z-x) d\hat{F}_n(x)}{h_{\hat{F}_n}(z)}.$$

**Solution (first step):** We introduce a right-continuous function  $B_{t, \hat{F}_n}$  such that

$$\frac{\int_{[0, z]} g(z - x) dB_{t, \hat{F}_n}(x)}{h_{\hat{F}_n}(z)} = a_{t, \hat{F}_n}(z), \quad \lim_{x \rightarrow \infty} B_{t, \hat{F}_n}(x) = 0,$$

where  $B_{t, \hat{F}_n}$  is no longer absolutely continuous w.r.t.  $\hat{F}_n$  and try again:

$$\begin{aligned} \int a_{t, \hat{F}_n}(z) dH_n(z) &= \int_{x \in [0, \infty)} \frac{\int_{x \in [0, z]} g(z - x) b_{t, \hat{F}_n}(x) d\hat{F}_n(x)}{h_{\hat{F}_n}(z)} dH_n(z) \\ &= \int_{x \in [0, \infty)} \int_{z \geq x} \frac{g(z - x)}{\hat{h}_n(z)} dH_n(z) dB_{t, \hat{F}_n}(x) \stackrel{?}{=} 0 \end{aligned}$$

Difficulty: characterization of MLE  $\hat{F}_n$  tells us that

$$\begin{aligned} \int_{z \geq x} \frac{g(z - x)}{\hat{h}_n(z)} dH_n(z) &\geq 1, \\ &= 1, \text{ if } x \text{ is a point of mass of } \hat{F}_n. \end{aligned}$$

**Solution (second step):** Introduce a function  $\bar{B}_{t, \hat{F}_n}$  that is constant on the same intervals as  $\hat{F}_n$  and equal to  $B_{t, \hat{F}_n}$  at points of mass of  $\hat{F}_n$ . Then:

$$\int_{x \in [0, \infty)} \int_{z \geq x} \frac{g(z - x)}{\hat{h}_n(z)} dH_n(z) d\bar{B}_{t, \hat{F}_n}(x) = \lim_{x \rightarrow \infty} \bar{B}_{t, \hat{F}_n}(x) = \lim_{x \rightarrow \infty} B_{t, \hat{F}_n}(x) = 0,$$

and, hopefully, the following difference will be “small”:

$$\int_{x \in [0, \infty)} \int_{z \geq x} \frac{g(z - x)}{\hat{h}_n(z)} dH_n(z) d\bar{B}_{t, \hat{F}_n}(x) - \int_{x \in [0, \infty)} \int_{z \geq x} \frac{g(z - x)}{\hat{h}_n(z)} dH_n(z) dB_{t, \hat{F}_n}(x).$$

Now:

$$\begin{aligned}
 K_t(\hat{F}_n) - K_t(F_0) &= \int_{[0,t)} \{g(t-x) - g(0)\} d\hat{F}_n(x) - \int_{[0,t)} \{g(t-x) - g(0)\} dF_0(x) \\
 &= \int a_{t,\hat{F}_n}(z) d(H_n - H_0)(z) + \int_{x \in [0,\infty)} \int_{z \geq x} \frac{g(z-x)}{\hat{h}_n(z)} dH_n(z) d(\bar{B}_{t,\hat{F}_n} - B_{t,\hat{F}_n})(x).
 \end{aligned}$$

**Solution (third step):** Prove that

$$\int a_{t,\hat{F}_n}(z) d(H_n - H_0)(z) = \int a_{t,F_0}(z) d(H_n - H_0)(z) + o_p(n^{-1/2}).$$

That's the general plan!

**Example**

We consider  $g(x) = 4(1 - x)^3 1_{[0,1]}(x)$  and  $F_0$  the Uniform(0, 1) distribution. Then we get the following equation in  $\phi_{t,F_0}(z) \stackrel{\text{def}}{=} \int_{x \in [0,z]} g(z - x) dB_{t,F_0}(x)$ :

$$\int_{z=x}^{x+1} a_{t,F_0}(z) g(z - x) dz = \int_{z=x}^{x+1} \frac{\phi_{t,F_0}(z)}{h_0(z)} g(z - x) dz = \{g(t - x) - g(0)\} 1_{[0,t]}(x) - K_t(F_0). \quad (4)$$

Writing  $\phi = \phi_{t,F_0}$  and  $B = B_{t,F_0}$  and  $a(x) = \phi(x)/h_0(x)$  we get by differentiating:

$$-4a(x) + 12 \int_{z=x}^{x+1} a(z)(1 + x - z)^2 dz = 12(1 - t + x)^2 \cdot 1_{[0,t]}(x), \quad x \neq t, \quad (5)$$

which leads to the following integral equation, using  $B(1) = 0$ ,

$$\begin{aligned} & \frac{B(x) - 3 \int_0^x (1 + u - x)^2 B(u) du}{h_0(x)} - 3 \int_{z=x}^1 \frac{B(z)(1 + x - z)^2}{h_0(z)} dz \\ & \quad + 9 \int_{u=0}^x B(u) \int_{z=x}^{1+u} \frac{(1 + u - z)^2 (1 + x - z)^2 dz}{h_0(z)} du \\ & \quad \quad \quad + 9 \int_{u=x}^1 B(u) \int_{z=u}^{1+x} \frac{(1 + u - z)^2 (1 + x - z)^2 dz}{h_0(z)} du \\ & = -\frac{3}{4}(1 - t + x)^2 \cdot 1_{[0,t]}(x), \quad x \neq t. \end{aligned}$$



We can also write this integral equation in the following form:

$$\begin{aligned}
 B(x) &- 3 \int_0^x B(u)(1+u-x)^2 du - 3h_0(x) \int_{u=x}^1 \frac{B(u)(1+x-u)^2}{h_0(u)} du \\
 &+ 9h_0(x) \int_{u=0}^x B(u) \int_{z=x}^{1+u} \frac{(1+u-z)^2(1+x-z)^2 dz}{h_0(z)} du \\
 &\quad + 9h_0(x) \int_{u=x}^1 B(u) \int_{z=u}^{1+x} \frac{(1+u-z)^2(1+x-z)^2 dz}{h_0(z)} du \\
 &= -\frac{3}{4}(1-t+x)^2 h_0(x) \cdot 1_{[0,t)}(x), \quad x \neq t.
 \end{aligned} \tag{6}$$

Introducing the notation

$$C_{t,F_0}(x) = C(x) = \frac{B(x)}{h_0(x)},$$

this can also be written as an integral equation in  $C(x)$ :

$$\begin{aligned}
 C(x) &- \frac{3}{h_0(x)} \int_0^x C(u)(1+u-x)^2 dH_0(u) - 3 \int_{u=x}^1 C(u)(1+x-u)^2 du \\
 &+ 9 \int_{u=0}^x C(u) \int_{z=x}^{1+u} \frac{(1+u-z)^2(1+x-z)^2 dz}{h_0(z)} dH_0(u) \\
 &\quad + 9 \int_{u=x}^1 C(u) \int_{z=u}^{1+x} \frac{(1+u-z)^2(1+x-z)^2 dz}{h_0(z)} dH_0(u) \\
 &= -\frac{3}{4}(1-t+x)^2 \cdot 1_{[0,t)}(x), \quad x \neq t.
 \end{aligned} \tag{7}$$

**Lemma 1.** *Let  $B_{t,F_0}$  and  $C_{t,F_0}$  be the solutions of the integral equation (6) and (7), respectively.*

- (i)  $B_{t,F_0}$  is non-positive, bounded and continuous on  $[0, 1]$  and  $B_{t,F_0}(0) = B_{t,F_0}(1) = 0$ . Moreover,  $B_{t,F_0}$  has a bounded derivative at each point  $x \in (0, 1) \setminus \{t\}$ , a jump of size  $\frac{3}{4}h_0(t)$  at  $t$ , a finite right derivative at  $x = 0$  and a left derivative, equal to zero, at  $x = 1$ .
- (ii)  $C_{t,F_0}$  is non-positive and bounded on  $(0, 1)$  with a bounded right limit at 0 and a left limit, equal to zero, at 1. Moreover,  $C_{t,F_0}$  has a bounded derivative at each point  $x \in (0, 1) \setminus \{t\}$ , a jump of size  $3/4$  at  $t$ , a finite right derivative at  $x = 0$  and a left derivative, equal to zero, at  $x = 1$ .
- (iii)  $a_{t,F_0}$  is bounded on  $(0, 2)$  with a bounded right limit at 0 and a left limit, equal to zero, at 2. Moreover,  $a_{t,F_0}$  has a bounded derivative at each point  $z \in (0, 2) \setminus \{t\}$ , a jump of size  $3/2$  at  $t$ , and finite right and left derivatives at  $z = 0$  and  $z = 2$ , respectively.

### The current status model

“Hidden space” variables are  $(T_i, X_i)$ ,  $T_i, X_i \in \mathbb{R}$ , observations are:  $(T_i, \Delta_i)$ .

$X_i$  is independent of  $T_i$ ,  $\Delta_i = 1_{\{X_i \leq T_i\}}$ . The  $X_i$  are (unobservable) “failure times”.

(Relevant part of) **Log likelihood** for the distribution function  $F$  of  $X_i$ :

$$\sum_{i=1}^n \{\Delta_i \log F(T_i) + (1 - \Delta_i) \log (1 - F(T_i))\}. \quad (8)$$

Define the empirical processes:

$$V_{n1}(t) = n^{-1} \sum_{T_i \leq t} \Delta_i, \quad V_{n2}(t) = n^{-1} \sum_{T_i \leq t} (1 - \Delta_i), \quad t \in \mathbb{R},$$

Then the log likelihood (8) for  $F$ , divided by  $n$ , can be written:

$$\int \log F(u) dV_{n1}(u) + \int \log\{1 - F(u)\} dV_{n2}(u). \quad (9)$$

## How can we determine the local behavior?

**Problem:** Unlike in  $\sqrt{n}$ -asymptotics, we do not have **global** convergence of the (rescaled) log likelihood process.

The situation is therefore fundamentally different from (1-dimensional) **right-censoring**, where for example the Kaplan-Meier estimator converges at  $\sqrt{n}$ -rate and maximizes a process which converges globally after rescaling.

But in the current situation we are lucky: **convex minorant interpretation of the MLE**.

**Proposition 1.** *Let  $H_n$  be the **greatest convex minorant** of the (so-called) **cusum diagram** (or **cumulative sum diagram**), consisting of the set of points*

$$\mathcal{P}_n = \left\{ \left( \mathbb{G}_n(t), V_{n1}(t) \right), t \in \mathbb{R} \right\}, \quad V_{n1}(t) = n^{-1} \sum_{i=1}^n \Delta_i 1_{(-\infty, t]}(T_i) \quad (10)$$

where  $\mathbb{G}_n(t) = n^{-1} \sum_{i=1}^n 1_{(-\infty, t]}(T_i)$  is the empirical distribution function of the observation times  $T_1, \dots, T_n$ .

Then  $\hat{F}_n$  is an MLE if and only if, at each observation point  $t = T_i$ ,  $\hat{F}_n(t)$  is the **left derivative of  $H_n$  at  $\mathbb{G}_n(t)$** .  $\hat{F}_n$  is uniquely determined at each observation point  $T_i$ .

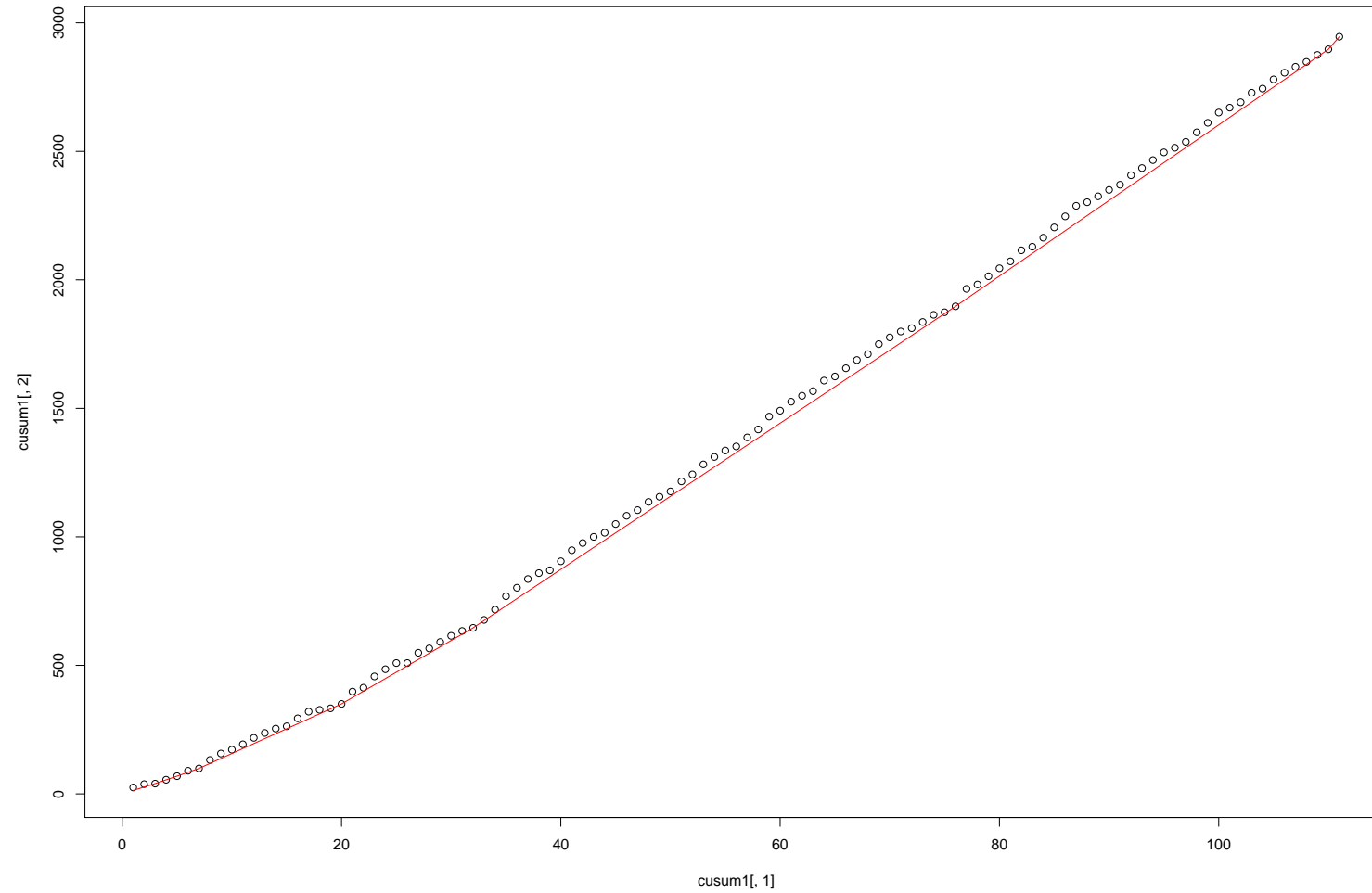


Figure 1: Cusum diagram. Simulation for  $n = 20$ . Observation df  $G$  of the  $T_i$  and df  $F$  of the  $X_i$  are both uniform.

## Local asymptotic behavior

Define

$$W_n(t) = n^{-1} \sum_{i=1}^n \{\Delta_i - F_0(T_i)\} \{1_{\{T_i \leq t\}} - 1_{\{T_i \leq t_0\}}\}, \quad t \in \mathbb{R}. \quad (11)$$

Let  $t_0$  be such that  $0 < F_0(t_0), G(t_0) < 1$ , and let  $F_0$  and  $G$  be continuously differentiable at  $t_0$ , with strictly positive derivatives  $f_0(t_0)$  and  $g(t_0)$ , respectively.

Then we have a “**Kim and Pollard (1990)-type lemma**”:

$$W_n(t) = O_p\left(n^{-2/3}\right) + o_p\left((t - t_0)^2\right), \quad \text{uniformly for } |t - t_0| \leq \delta. \quad (12)$$

After rescaling, the MLE  $\hat{F}_n$  is the **slope of the convex minorant** of the process

$$U_n(t) \stackrel{\text{def}}{=} n^{2/3} W_n\left(t_0 + n^{-1/3}t\right) + n^{-1/3} \sum_{i=1}^n \{F_0(T_i) - F_0(t_0)\} \{1_{\{T_i \leq t\}} - 1_{\{T_i \leq t_0\}}\}, \quad t \in \mathbb{R},$$

which converges to two-sided (scaled) Brownian motion with a parabolic drift. We can **localize**, due to the fact that, for large  $|t|$ , the drift in the process  $U_n$  is dominated by the **parabolic** drift of

$$n^{-1/3} \sum_{i=1}^n \{F_0(T_i) - F_0(t_0)\} \{1_{\{T_i \leq t\}} - 1_{\{T_i \leq t_0\}}\} \sim \frac{1}{2} f_0(t_0) t^2, \quad n \rightarrow \infty.$$

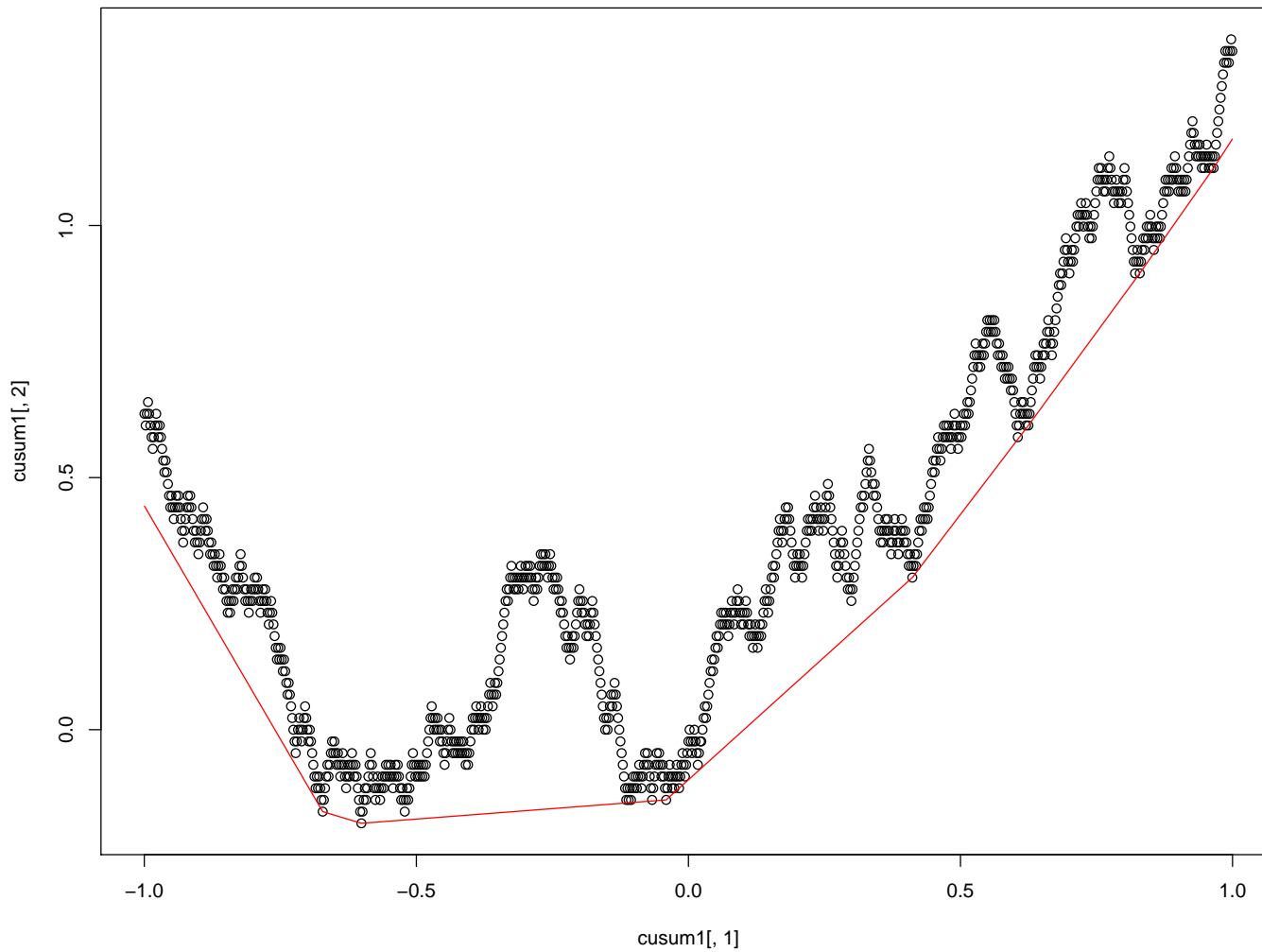


Figure 2: **Locally rescaled cusum process**  $(n^{1/3}(\mathbb{G}_n(t_0 + n^{1/3}t) - \mathbb{G}_n(t_0)), U_n(t))$ , **with convex minorant**, for  $n^{1/3}|\mathbb{G}_n(t_0 + n^{1/3}t) - \mathbb{G}_n(t_0)| \leq 1$  **and**  $t_0 = 0.5$ . Simulation with  $n = 10,000$ . Observation df  $G$  of the  $T_i$  and df  $F$  of the  $X_i$  are both **uniform**.

## Local limit distribution of MLE $\hat{F}_n$

Using the “Kim and Pollard (1990)-type lemma”:

$$W_n(t) = O_p\left(n^{-2/3}\right) + o_p\left((t - t_0)^2\right), \text{ uniformly for } |t - t_0| \leq \delta,$$

we can “localize” the convex minorant and hence its derivative process, yielding the MLE  $\hat{F}_n$ .

Sketch of derivation of local limit distribution:

### 1. The localized cusum diagram

$$\left(n^{1/3} \left(\mathbb{G}_n(t_0 + n^{1/3}t) - t_0\right), U_n(t)\right)$$

converges in distribution to the **Brownian motion cusum diagram**:

$$(g(t_0)t, U(t)), \text{ where } U(t) = \sqrt{g(t_0)F_0(t_0)\{1 - F_0(t_0)\}} W(t) + \frac{1}{2}f_0(t_0)g(t_0)t^2, t \in \mathbb{R},$$

and where  $W$  is two-sided Brownian motion.

2. Continuous mapping theorem: **Convex minorant of localized cusum diagram** converges in distribution to **convex minorant of Brownian cusum diagram**.
3. **Left-derivative of convex minorant of localized cusum diagram** converges in distribution (in Skohorod topology) to **left-derivative of convex minorant of Brownian cusum diagram** and  $n^{1/3}\{\hat{F}_n(t_0) - F_0(t_0)\}$  **is left-derivative of convex minorant of localized cusum diagram at zero**.



## Competing risk model

Generalization of the current status model to the situation where there are more failure causes.

Hidden space variables are  $(T_i, X_i, Y_i)$ ,  $Y_i$  is the failure cause.

Observations are:  $(T_i, \Delta_{i1}, \dots, \Delta_{iK})$ ,  $\Delta_{ik} = 1_{\{X_i \leq T_i, Y_i = k\}}$ . Define:

$$V_{nk}(t) = n^{-1} \sum_{i=1}^n \Delta_{ik} 1_{(-\infty, t]}(T_i), \quad V_{n, K+1}(t) = n^{-1} \sum_{i=1}^n (1 - \Delta_{i+}) 1_{(-\infty, t]}(T_i), \quad \Delta_{i+} = \sum_{k=1}^K \Delta_{ik},$$

We want to estimate the **subdistribution functions**  $F_{0k}$ :

$$F_{0k}(t) = P \{X \leq t, Y = k\}, \quad k = 1, \dots, K.$$

The (relevant part of the) **log likelihood** for  $F = (F_1, \dots, F_K)$ , divided by  $n$ , is:

$$\sum_{k=1}^K \int \log F_k(u) dV_{nk}(u) + \int \log \{1 - F_+(u)\} dV_{n, K+1}(u), \quad F_+ = \sum_{k=1}^K F_k.$$

The **MLE** (maximum likelihood estimator)  $\hat{F}_n = (\hat{F}_{n1}, \dots, \hat{F}_{nK})$  can only be computed **iteratively**.

**No direct convex minorant interpretation**, as with the MLE for current status data.

## Self-induced characterization

The MLE  $\hat{F}_{nk}$  can be characterized as the left derivative of the greatest convex minorant of the **self-induced cusum diagram**

$$\mathcal{P}_{nk} = \left\{ (G_{\hat{F}_{n+}}(t), V_{nk}(t)), t \in \mathbb{R} \right\}, \quad V_{nk}(t) = n^{-1} \sum_{i=1}^n \Delta_{ik} 1_{(-\infty, t]}(T_i), \quad (13)$$

for  $k = 1, \dots, K$ , where

$$G_{\hat{F}_{n+}}(t) = n^{-1} \sum_{i=1}^n \frac{1 - \Delta_{i+}}{1 - \hat{F}_{n+}(T_i)} 1_{(-\infty, t]}(T_i), \quad t < T_{(n)}.$$

Compare with ordinary current status, where  $\hat{F}_n$  is the left derivative of the greatest convex minorant of the **(not self-induced) cusum diagram**

$$\mathcal{P}_n = \left\{ (\mathbb{G}_n(t), V_{n1}(t)), t \in \mathbb{R} \right\}, \quad V_{n1}(t) = n^{-1} \sum_{i=1}^n \Delta_i 1_{(-\infty, t]}(T_i). \quad (14)$$

Note:

$$G_{\hat{F}_{n+}}(t) = \mathbb{G}_n(t) + n^{-1} \sum_{i=1}^n \frac{\hat{F}_{n+}(T_i) - \Delta_{i+}}{1 - \hat{F}_{n+}(T_i)} 1_{(-\infty, t]}(T_i), \quad t < T_{(n)}.$$

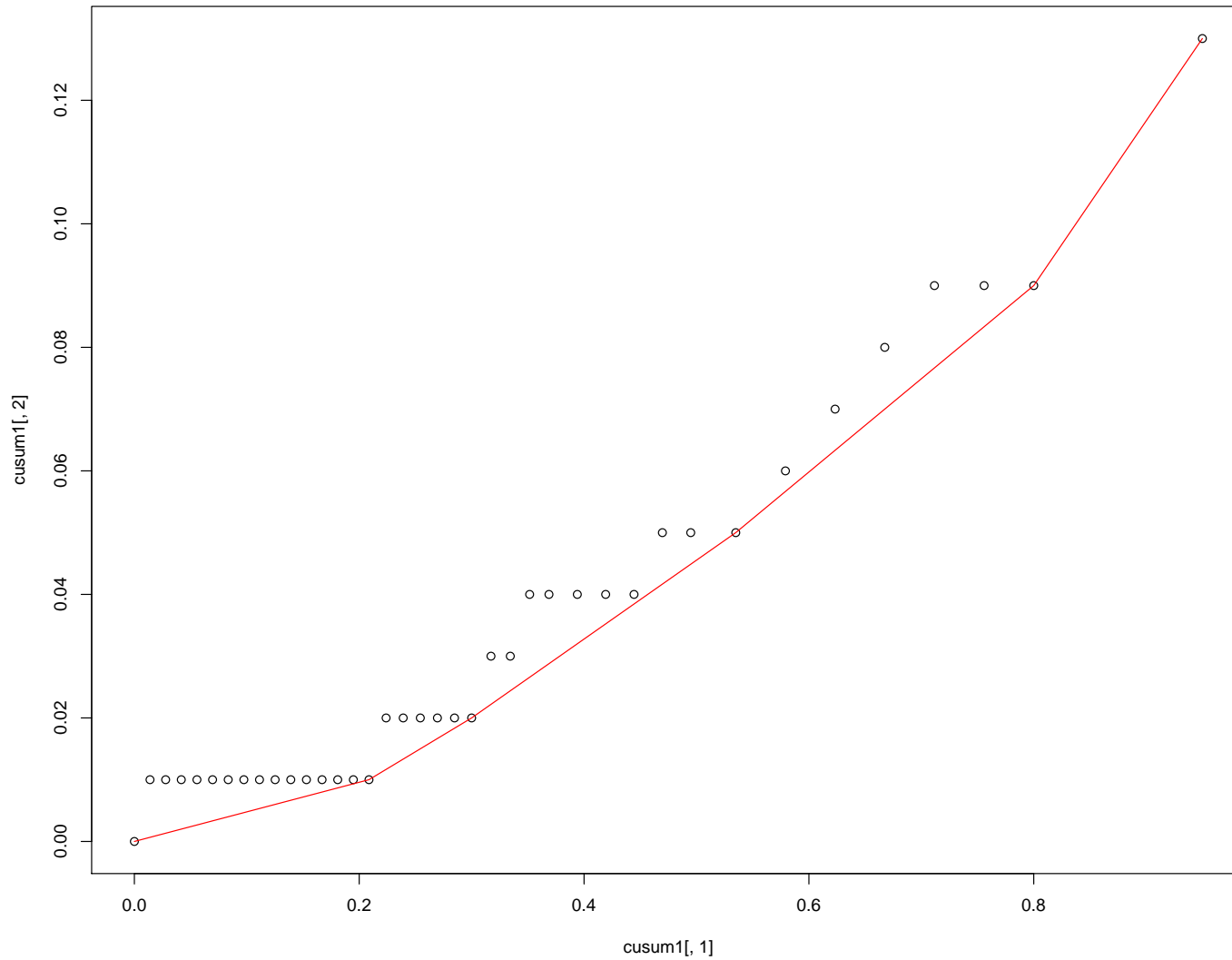


Figure 3: **Cusum diagram**  $\{(G_{\hat{F}_{n+}}(t), V_{n1}(t)), t \in \mathbb{R}\}$ . Simulation for  $n = 100$ ,  $K = 2$ .  $F_{0k}(t) = (k/3)\{1 - e^{-kt}\}$ ,  $T \sim \text{Unif}(0, 1.5)$ .

## Some history of work on the local rate

**End of 2004:** The following fact was proved.

**Lemma 2.** *Let  $\hat{F}_n = (\hat{F}_{n1}, \dots, \hat{F}_{nK})$  be the MLE of  $F_0 = (F_{01}, \dots, F_{0K})$ , and let  $\hat{F}_{n+} = \sum_{k=1}^K \hat{F}_{nk}$  and, similarly,  $F_{0+} = \sum_{k=1}^K F_{0k}$ . Moreover, let, for a  $\delta \in (0, 1)$ ,  $[t_0 - \delta, t_0 + \delta]$  be an interval on which the components  $F_{0k}$  have continuous derivatives staying away from zero. Then there exists for each  $\varepsilon > 0$  and  $M > 0$  an  $M_1 > 0$  so that*

$$\mathbb{P} \left\{ \sup_{t \in [-M, M]} n^{1/3} \left| \hat{F}_{n+}(t_0 + n^{-1/3}t) - F_{0+}(t_0) \right| > M_1 \right\} < \varepsilon, \quad k = 1, \dots, K.$$

**This is not enough!** To get localization of the  $\hat{F}_{nk}$ , we need that for  $t$  outside a neighborhood of order  $O(n^{-1/3})$  of  $t_0$  we can replace  $G_{\hat{F}_{n+}}(t)$  by

$$G_{F_{0+}}(t) \stackrel{\text{def}}{=} \mathbb{G}_n(t) + n^{-1} \sum_{i=1}^n \frac{F_{0+}(T_i) - \Delta_{i+}}{1 - F_{0+}(T_i)} 1_{(-\infty, t]}(T_i),$$

up to terms of order  $o_p((t - t_0)^2)$  in the self-induced cusum diagram:

$$\mathcal{P}_{nk} = \left\{ (G_{\hat{F}_{n+}}(t), V_{nk}(t)), t \in \mathbb{R} \right\}, \quad V_{nk}(t) = n^{-1} \sum_{i=1}^n \Delta_{ik} 1_{(-\infty, t]}(T_i), \quad (15)$$

removing the self-inducedness of the coordinate  $G_{\hat{F}_{n+}}$ .

**Solution in September 2005: strengthening of Kim-Pollard-type lemma.**

**Lemma 3. (Global to local lemma for  $\hat{F}_{n+}$ )** Let  $\hat{F}_n$  be the MLE and let, for a  $\delta \in (0, 1)$ ,  $[t_0 - 2\sqrt{\delta}, t_0 + 2\sqrt{\delta}]$  be an interval on which the components  $F_{0k}$  have continuous derivatives staying away from zero. Then, for all  $t \in [t_0 - \delta, t_0 + \delta]$  we have:

$$\int_{\{|u-t_0| < |t-t_0|\}} \frac{|\hat{F}_{n+}(u) - F_{0+}(u)|}{1 - \hat{F}_{n+}(u)} d\mathbb{G}_n(u) = n^{-1/6} O_p \left( n^{-1/2} \vee |t - t_0|^{3/2} \right), \text{ uniformly in } t \in [t_0 - \delta, t_0 + \delta].$$

**Note:**  $n^{-1/6} O_p \left( n^{-1/2} \vee |t - t_0|^{3/2} \right) = O_p(n^{-2/3})$  if  $|t - t_0| = O_p(n^{-1/3})$ .

**Corollary 1. (Tightness of  $n^{1/3}\{\hat{F}_n(t_0 + n^{-1/3}t) - F_0(t_0)\}$ )** Let the conditions of Lemma 3 be satisfied. Then:

(i) (Replacement of  $G_{\hat{F}_{n+}}$  by  $G_{F_{0+}}$ )

$$G_{\hat{F}_{n+}}(t) = G_{F_{0+}}(t) + n^{-1/6} O_p \left( n^{-1/2} \vee |t - t_0|^{3/2} \right), \text{ uniformly in } t \in [t_0 - \delta, t_0 + \delta].$$

(ii) For each  $\varepsilon > 0$  and  $M > 0$  there exists an  $M_1 > 0$  so that

$$\mathbb{P} \left\{ \sup_{t \in [-M, M]} n^{1/3} \left| \hat{F}_{nk}(t_0 + n^{-1/3}t) - F_{0k}(t_0) \right| > M_1 \right\} < \varepsilon, \quad k = 1, \dots, K.$$

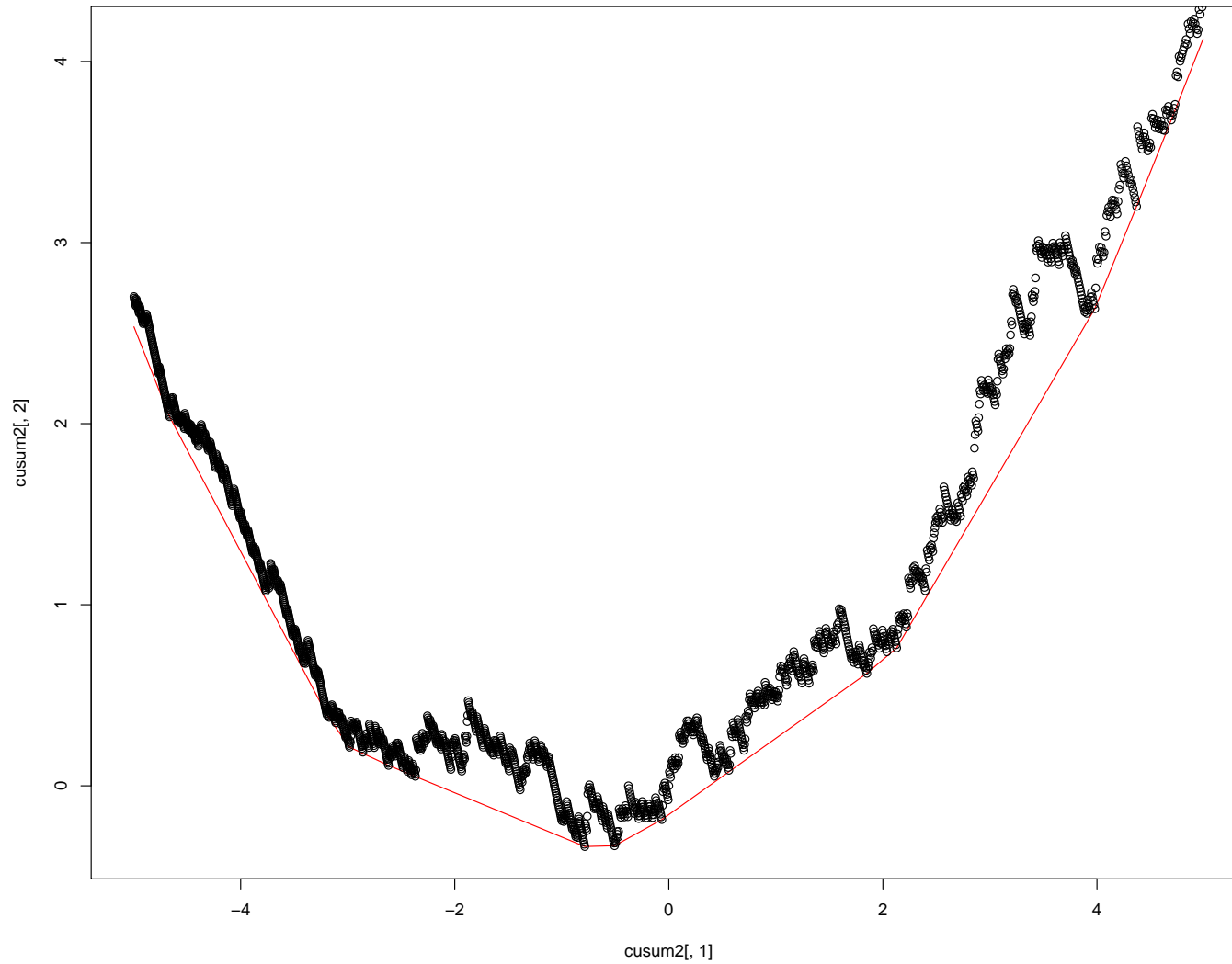


Figure 4: Localized cusum diagram  $\left\{ \left( n^{1/3} (G_{\hat{F}_{n+}}(t) - G_{\hat{F}_{n+}}(t_0)), n^{2/3} \left\{ V_{n1}(t) - V_{n1}(t_0) - \int_{t_0}^t F_{01}(t_0) dG_{\hat{F}_{n+}}(u) \right\} \right), t \in \mathbb{R} \right\}$ ,  $t_0 = 0.5, n = 10,000$ . **Red curve:**  $n^{2/3} \int_{t_0}^t \{ \hat{F}_{n1}(t_0) - F_{01}(u) \} dG_{\hat{F}_{n+}}(u)$ . If  $t < t_0$ :  $\int_{t_0}^t \{ \hat{F}_{n1}(u) - F_{01}(t_0) \} dG_{\hat{F}_{n+}}(u) \stackrel{\text{def}}{=} - \int_{[t, t_0]} \{ \hat{F}_{n1}(u) - F_{01}(t_0) \} dG_{\hat{F}_{n+}}(u)$ .

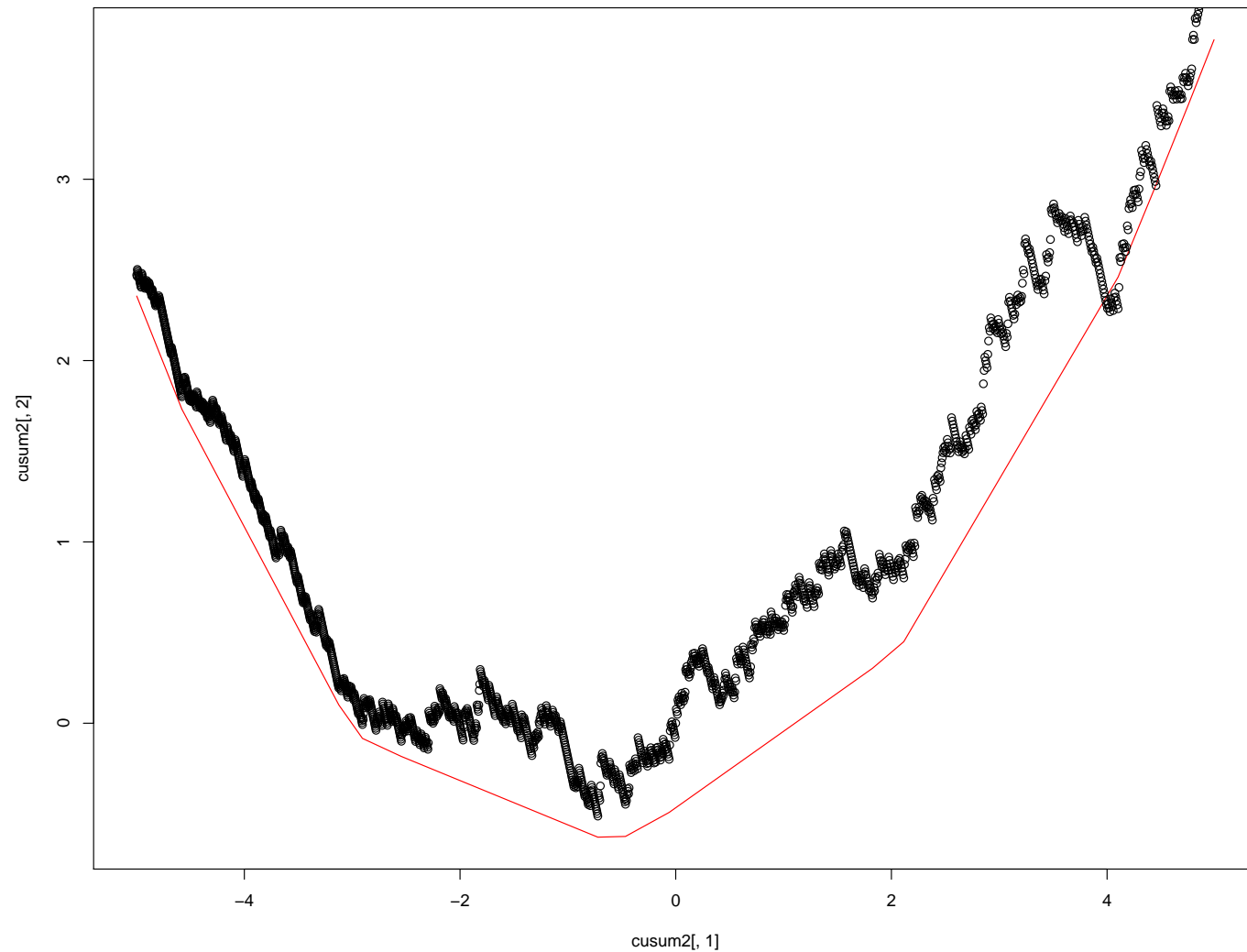


Figure 5: Localized cusum diagram with  $G_{\hat{F}_{n+}}$  replaced by  $G_{F_{0+}}$ ,  $t_0 = 0.5, n = 10,000$ . **Red curve:**  $n^{2/3} \int_{t_0}^t \{\hat{F}_{n1}(u) - F_{01}(t_0)\} dG_{F_{0+}}(u)$ . If  $t < t_0$ :  $\int_{t_0}^t \{\hat{F}_{n1}(u) - F_{01}(t_0)\} dG_{F_{0+}}(u) \stackrel{\text{def}}{=} - \int_{[t, t_0]} \{\hat{F}_{n1}(u) - F_{01}(t_0)\} dG_{F_{0+}}(u)$ .

## Summary of the global to local argument

The MLE maximizes a **global** criterion. To extract the **local** limit behavior from this, we have to use some kind of characterization of the solution, for example a convex duality criterion.

1. In the case of **simple current status data**, this leads to a convex minorant characterization, which can be used for the determining the local behavior of the MLE.
2. In the case of **competing risk with current status data**, this leads to a **self-induced** convex minorant characterization, involving the sum  $\hat{F}_{n+}$  of the individual MLE estimators  $\hat{F}_{nk}$  for the several subdistribution functions  $F_{0k}$ . To get localization of the  $\hat{F}_{nk}$ , we need that for  $t$  outside a neighborhood of order  $O(n^{-1/3})$  of  $t_0$  we can replace  $G_{\hat{F}_{n+}}(t)$  by

$$G_{F_{0+}}(t) \stackrel{\text{def}}{=} \mathbb{G}_n(t) + n^{-1} \sum_{i=1}^n \frac{F_{0+}(T_i) - \Delta_{i+}}{1 - F_{0+}(T_i)} 1_{(-\infty, t]}(T_i),$$

up to terms of order  $o_p((t - t_0)^2)$  in the **self-induced cusum diagram**:

$$\mathcal{P}_{nk} = \left\{ (G_{\hat{F}_{n+}}(t), V_{nk}(t)), t \in \mathbb{R} \right\}, V_{nk}(t) = n^{-1} \sum_{i=1}^n \Delta_{ik} 1_{(-\infty, t]}(T_i), \quad (16)$$

to get rid of the self-inducedness of the coordinate  $G_{\hat{F}_{n+}}$  in the tightness argument.

This is accomplished by the **global to local lemma**.



### Limit distribution for competing risk

- First prove uniqueness of the limiting process, using tightness argument (Hardest part!)
- Localize characterization of limit process.
- Take subsequences of localized processes, based on a samples of size  $n$ , on  $[-m, m]$ . By tightness (using local rate result) there is a further subsequence that converges to some limit. Using a diagonal argument, it follows that there is a limit on  $\mathbb{R}$ . Here we go from local to global!
- By the continuous mapping theorem the limit must satisfy the limit characterization on  $[-m, m]$  for each  $m \in \mathbb{N}$ .
- Letting  $m \rightarrow \infty$  gives existence of the limiting process (almost for free!)
- By uniqueness of the limiting process, all subsequences converge to the same limit

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